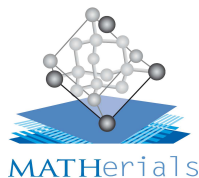


Simulation moléculaire : modèles, calcul d'énergie libre et aspects dynamiques

3- Echantillonnage de dynamiques métastables

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Introduction

Remember the dynamics:

- *Langevin* dynamics:

$$\begin{cases} d\mathbf{X}_t = M^{-1}\mathbf{P}_t dt \\ d\mathbf{P}_t = -\nabla V(\mathbf{X}_t) dt - \gamma M^{-1}\mathbf{P}_t dt + \sqrt{2\gamma\beta^{-1}}d\mathbf{W}_t \end{cases}$$

where $\gamma > 0$ and $\beta = (k_B T)^{-1}$.

- *overdamped Langevin* (or gradient) dynamics:

$$d\mathbf{X}_t = -\nabla V(\mathbf{X}_t) dt + \sqrt{2\beta^{-1}}d\mathbf{W}_t.$$

Introduction

These dynamics are used to compute macroscopic quantities:

- (i) Thermodynamic quantities (averages wrt ν of some observables): stress, heat capacity, free energy,...

$$\mathbb{E}_\nu(\varphi(\mathbf{X})) = \int_{\mathbb{R}^d} \varphi(\mathbf{x}) \nu(d\mathbf{x}) \simeq \frac{1}{T} \int_0^T \varphi(\mathbf{X}_t) dt.$$

- (ii) Dynamical quantities (averages over trajectories): diffusion coefficients, viscosity, transition rates,...

$$\mathbb{E}(\mathcal{F}((\mathbf{X}_t)_{t \geq 0})) \simeq \frac{1}{M} \sum_{m=1}^M \mathcal{F}((\mathbf{X}_t^m)_{t \geq 0}).$$

Difficulties: (i) high-dimensional problem ($N \gg 1$); (ii) \mathbf{X}_t is a metastable process and ν is a multimodal measure.

Introduction

For computing thermodynamics quantities, there is a clear classification of available methods, and the difficulties are now well understood (in particular for free energy computations, see for example [TL, Rousset, Stoltz, 2010]). On the opposite, **computing efficiently dynamical quantities remains a challenge.**

Outline of this part:

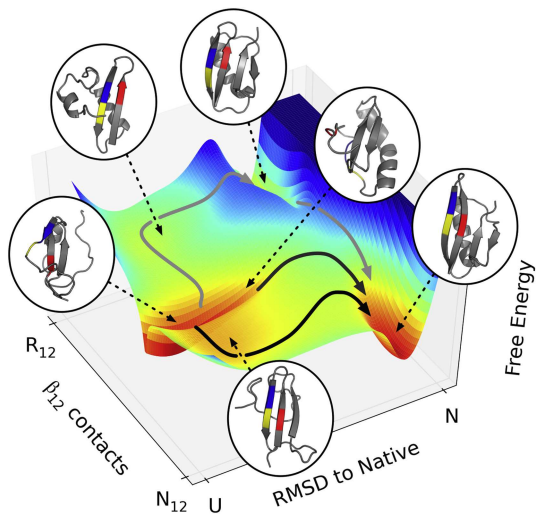
1. **Accelerated dynamics**: These methods have been proposed by A.F. Voter to generate efficiently metastable dynamics. *Mathematical tool: Quasi Stationary Distributions.*
2. **Adaptive Multilevel Splitting** methods: Towards efficient sampling of reactive paths. *Rare event simulation.*

Underlying question: how to properly define and quantify metastability ? Various answers: (i) rate of convergence to equilibrium; (ii) exit time from metastable states; (iii) decorrelation time; (iv) asymptotic variance of estimators.

Accelerated dynamics

- 1- From Langevin to kMC
- 2- From theory to algorithms

1- From Langevin to kinetic Monte Carlo



C.R. Schwantes, D. Shukla, V.S.Pande, Biophysical Journal, vol. 110, 2016

Two models for dynamics

The basic modeling ingredient in molecular dynamics: a **potential function** V which associates to a configuration $\mathbf{x} = (\mathbf{x}_1, \dots, \mathbf{x}_{N_{atom}}) \in \mathbb{R}^{3N_{atom}}$ an energy $V(\mathbf{x}) \in \mathbb{R}$.

From V , two kinds of dynamics are considered:

- Langevin and over-damped Langevin dynamics: Markov processes with values in continuous state space ;
- kinetic Monte Carlo model or Markov state model (first order kinetics): Markov processes with values in discrete state space (jump Markov process).

Question: **Can a mathematically rigorous link be made between these two kinds of models ?**

Langevin and over-damped Langevin dynamics

Let us introduce the inverse temperature: $\beta^{-1} = k_B T$.

The Langevin dynamic writes:

$$\begin{cases} d\mathbf{X}_t = M^{-1}\mathbf{P}_t dt, \\ d\mathbf{P}_t = -\nabla V(\mathbf{X}_t) dt - \gamma M^{-1}\mathbf{P}_t dt + \sqrt{2\gamma\beta^{-1}} d\mathbf{W}_t. \end{cases}$$

In the following, we focus on the *over-damped Langevin* dynamics

$$d\mathbf{X}_t = -\nabla V(\mathbf{X}_t) dt + \sqrt{2\beta^{-1}} d\mathbf{W}_t.$$

These dynamics are both ergodic wrt the canonical measure:

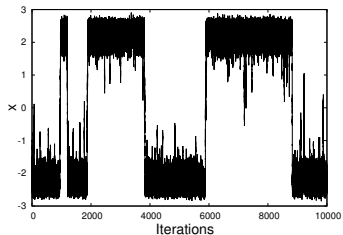
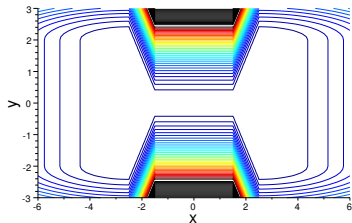
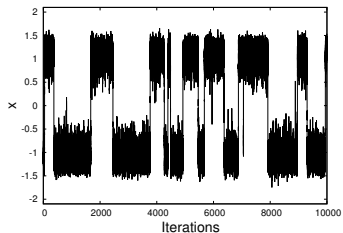
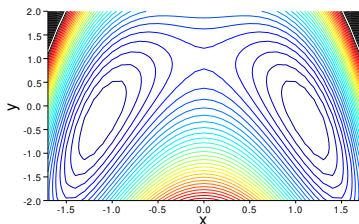
$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t \varphi(\mathbf{X}_s) ds = \int \varphi d\nu$ where

$$\nu(dx) = Z^{-1} \exp(-\beta V(x)) dx.$$

Main practical challenge: these dynamics are [metastable](#).

Metastability: energetic and entropic barriers

A two-dimensional schematic picture



-
- Slow convergence of trajectorial averages
 - Transitions between metastable states are **rare events**

Metastability: a toy example

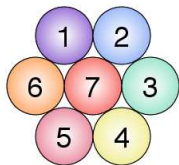
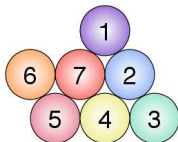
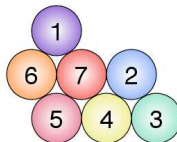
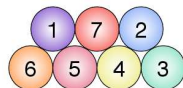
(a) $V = -12.53$ (b) $V = -11.50$ (c) $V = -11.48$ (d) $V = -11.40$

Figure: Low energy conformations of the 7 atoms Lennard-Jones cluster.

→ simulation

The exit event

Let us consider the overdamped Langevin dynamics:

$$d\mathbf{X}_t = -\nabla V(\mathbf{X}_t) dt + \sqrt{2\beta^{-1}} d\mathbf{W}_t$$

and let assume that we are given an ensemble of subsets of \mathbb{R}^d (states). Let us consider one of them: $\mathcal{S} \subset \mathbb{R}^d$. The **exit event** from \mathcal{S} is given by

$$(\tau_{\mathcal{S}}, \mathbf{X}_{\tau_{\mathcal{S}}})$$

where $\tau_{\mathcal{S}} = \inf\{t > 0, \mathbf{X}_t \notin \mathcal{S}\}$.

Objective: build a jump Markov model to simulate the exit event $(\tau_{\mathcal{S}}, \mathbf{X}_{\tau_{\mathcal{S}}})$.

This is useful theoretically (justification of Markov state models and Eyring-Kramers laws) and numerically (accelerated dynamics *à la* Voter).

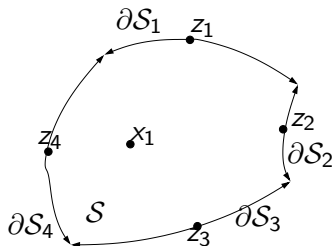
Kinetic Monte Carlo

Kinetic Monte Carlo (or Markov state) models are built as follows:

- define exit regions from \mathcal{S} : $\partial\mathcal{S} = \cup_{j=1}^J \partial\mathcal{S}_j$
- associate a rate k_j with an exit through $\partial\mathcal{S}_j$

and then (jump Markov model)

- the exit time τ_S^{kMC} is exponentially distributed with parameter $\sum_{j=1}^J k_j$
- the exit region is I_S^{kMC} with law $\mathbb{P}(I_S^{kMC} = i) = \frac{k_i}{\sum_{j=1}^J k_j}$
- I_S^{kMC} and τ_S^{kMC} are independent random variables

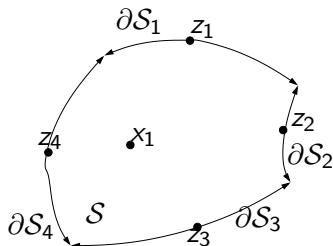


Eyring-Kramers laws

Formulas for transition rates. Let us introduce the local minima $(z_j)_{j=1,\dots,J}$ of V on ∂S , and associated exit regions ∂S_i . The parameters k_j are computed using the Eyring-Kramers formula (Harmonic Transition State Theory):

$$k_j^{HTST} = \nu_j e^{-\beta[V(z_j) - V(x_1)]}$$

where ν_j is an explicit prefactor and $x_1 = \arg \min_S V$.



A theoretical question

Question: can we relate the exit event $(\tau_S, \mathbf{X}_{\tau_S})$ for the original dynamics with the exit event $(\tau_S^{kMC}, I_S^{kMC})$ for the jump Markov process?

Two steps:

- Introduce the Quasi-Stationary Distribution
- Consider the small temperature regime $\beta \rightarrow \infty$
(semi-classical limit)

Step 1: The Quasi-Stationary Distribution

Definition of the QSD: Let \mathbf{X}_0 start in the state \mathcal{S} . Then there exists a probability distribution ν with support \mathcal{S} such that

$$\lim_{t \rightarrow \infty} \mathcal{L}(\mathbf{X}_t | \tau_{\mathcal{S}} > t) = \nu$$

where $\tau_{\mathcal{S}}$ is the first exit time from \mathcal{S} .

Remark: Quantitative definition of a metastable exit:
exit time \gg local equilibration time

Fundamental property of the QSD: Starting from ν :

- the first exit time $\tau_{\mathcal{S}}$ is **exponentially distributed** ;
- and $\tau_{\mathcal{S}}$ is **independent of the first hitting point** $\mathbf{X}_{\tau_{\mathcal{S}}}$.

Consequence: Starting from ν , the exit event from \mathcal{S} can be **exactly** written as one jump of a kinetic Monte Carlo model with rates

$$k_i = \frac{\mathbb{P}^{\nu}(\mathbf{X}_{\tau_{\mathcal{S}}} \in \partial \mathcal{S}_i)}{\mathbb{E}^{\nu}(\tau_{\mathcal{S}})}.$$

Step 2: The small temperature regime

Moreover, one has **explicit formulas for $\mathbb{E}(\tau_S)$ and the distribution of \mathbf{X}_{τ_S}** . Let us introduce the first eigenstate (λ_1, u_1) of the Fokker-Planck operator associated with the dynamics with Dirichlet boundary conditions on $\partial\mathcal{S}$:

$$\begin{cases} \operatorname{div}(\nabla V u_1) + \beta^{-1} \Delta u_1 = -\lambda_1 u_1 \text{ on } \mathcal{S}, \\ u_1 = 0 \text{ on } \partial\mathcal{S}. \end{cases}$$

Then, $\nu = \frac{u_1(\mathbf{x}) d\mathbf{x}}{\int_{\mathcal{S}} u_1}$,

$$\mathbb{E}^\nu(\tau_S) = \frac{1}{\lambda_1}$$

and

$$\mathbb{P}^\nu(\mathbf{X}_{\tau_S} \in \partial\mathcal{S}_i) = -\frac{\int_{\partial\mathcal{S}_i} \partial_n u_1 d\sigma}{\beta \lambda_1 \int_{\mathcal{S}} u_1(\mathbf{x}) d\mathbf{x}}.$$

Thus, $k_i = -\frac{\int_{\partial\mathcal{S}_i} \partial_n u_1 d\sigma}{\beta \lambda_1 \int_{\mathcal{S}} u_1(\mathbf{x}) d\mathbf{x}}$.

Can we then show that $k_i \simeq k_i^{HTST}$?

Justifying Eyring-Kramers laws

Theorem [Di Gesu, TL, Le Peutrec, Nectoux, 2019]

Under some geometric assumptions, starting from the QSD, in the limit $\beta \rightarrow \infty$, the exit rates are

$$k_i = \tilde{\nu}_i^{OL} e^{-\beta[V(z_i) - V(x_1)]} (1 + O(\beta^{-1}))$$

where

$$\tilde{\nu}_i^{OL} = \sqrt{\frac{\beta}{2\pi}} \partial_n V(z_i) \frac{\sqrt{\det(\nabla^2 V)(x_1)}}{\sqrt{\det(\nabla^2 V|_{\partial S})(z_i)}}.$$

Assumptions (1/2)

- \mathcal{S} is an open bounded smooth domain in \mathbb{R}^d .
- $V : \overline{\mathcal{S}} \rightarrow \mathbb{R}$ is a Morse function with a single critical point x_1 .
Moreover, $x_1 \in \mathcal{S}$ and $V(x_1) = \min_{\overline{\mathcal{S}}} V$.
- $\partial_n V > 0$ on $\partial\mathcal{S}$ and $V|_{\partial\mathcal{S}}$ is a Morse function with local minima reached at z_1, \dots, z_J with $V(z_1) < \dots < V(z_J)$.
- $V(z_1) - V(x_1) > V(z_J) - V(z_1)$
- $\forall i \in \{1, \dots, J\}$, consider B_{z_i} the basin of attraction of z_i for the dynamics $\dot{x} = -\nabla_T V(x)$ and assume that

$$\inf_{z \in B_{z_i}^c} d_a(z, z_i) > V(z_J) - V(z_1)$$

Assumptions (2/2)

Here, d_a is the Agmon distance:

$$d_a(x, y) = \inf_{\gamma} \int_0^1 g(\gamma(t)) |\gamma'(t)| dt$$

where $g = \begin{cases} |\nabla V| & \text{in } \mathcal{S} \\ |\nabla_T V| & \text{in } \partial\mathcal{S} \end{cases}$, and the infimum is over all piecewise \mathcal{C}^1 paths $\gamma : [0, 1] \rightarrow \overline{\mathcal{S}}$ such that $\gamma(0) = x$ and $\gamma(1) = y$.

Numerical tests indicate that the assumption

$$\forall i \in \{1, \dots, J\}, \inf_{z \in B_{z_i}^c} d_a(z, z_i) > V(z_i) - V(z_1)$$

seems indeed necessary to get the expected results.

Sketch of the proof (1/3)

The difficult part is to find an approximation for

$$\int_{\partial S_i} \partial_n u_1 = \int_{\partial S_i} \partial_n v_1 e^{-\beta V}, \text{ where } v_1 = u_1 e^{\beta V}.$$

We have

$$\begin{cases} L^{(0)} v_1 = -\lambda_1 v_1 \text{ on } W, \\ v_1 = 0 \text{ on } \partial W, \end{cases}$$

where $L^{(0)} = \beta^{-1} \Delta - \nabla V \cdot \nabla$ is a self adjoint operator on $L^2(e^{-\beta V})$. We are interested in $\nabla v_1 \cdot n$, and ∇v_1 satisfies

$$\begin{cases} L^{(1)} \nabla v_1 = -\lambda_1 \nabla v_1 \text{ on } W, \\ \nabla_T v_1 = 0 \text{ on } \partial W, \\ (\beta^{-1} \operatorname{div} - \nabla V \cdot) \nabla v_1 = 0 \text{ on } \partial W, \end{cases}$$

where

$$L^{(1)} = \beta^{-1} \Delta - \nabla V \cdot \nabla - \operatorname{Hess}(V).$$

Therefore ∇v_1 is an eigenvector (eigen-1-form) of $-L^{(1)}$ associated with the small eigenvalue λ_1 .

Sketch of the proof (2/3)

Let $\Pi^{(\rho)} = 1_{[0, \beta^{-3/2}]}(-L^{(\rho)})$ be the spectral projection operator on small eigenvalues. We know [Helffer, Sjöstrand] that, for β large, $\dim(\text{Ran}\Pi^{(0)}) = 1$ and $\dim(\text{Ran}\Pi^{(1)}) = J$:

$$\text{Ran}\Pi^{(0)} = \text{Span}(v_1)$$

$$\text{Ran}\Pi^{(1)} = \text{Span}(\psi_1, \dots, \psi_J).$$

Since $\nabla v_1 \in \text{Ran}\Pi^{(1)}$,

$$\int_{\partial S_i} \partial_n v_1 e^{-\beta V} = \sum_{j=1}^J \langle \nabla v_1, \psi_j \rangle_{L^2(e^{-\beta V})} \int_{\partial S_i} \psi_j \cdot n e^{-\beta V}.$$

The idea is now to build so-called **quasi-modes** which approximate the eigenvectors of $L^{(0)}$ and $L^{(1)}$ associated with small eigenvalues in the regime $\beta \rightarrow \infty$, in order to approximate the terms in the sum.

Sketch of the proof (3/3)

- $\text{Ran}\Pi^{(0)}$: an approximation of v_1 is given by

$$\tilde{v} = Z^{-1}\chi_{W'}$$

where $W' \subset\subset W$.

- $\text{Ran}\Pi^{(1)}$: an approximation of $\text{Ran}\Pi^{(1)}$ is $\text{Span}(\tilde{\psi}_1, \dots, \tilde{\psi}_J)$ where $(\tilde{\psi}_i)_{1 \leq i \leq J}$ are solutions to auxiliary eigenvalue problems, attached to the local minima $(z_i)_{1 \leq i \leq J}$.

Two tools:

- Agmon estimates (the support of $\tilde{\psi}_i$ is essentially in a neighborhood of z_i):

$$\exists N > 0, \|e^{\beta d_a(z_i, \cdot)/2} \tilde{\psi}_i\|_{H^1(e^{-\beta V})} = O(\beta^N).$$

- WKB approximations:

$$\exists N > 0, \tilde{\psi}_i \simeq Z_i^{-1} d(e^{\beta V/2} e^{-\beta d_a(z_i, \cdot)/2}) \beta^p.$$

Generalizations and perspectives

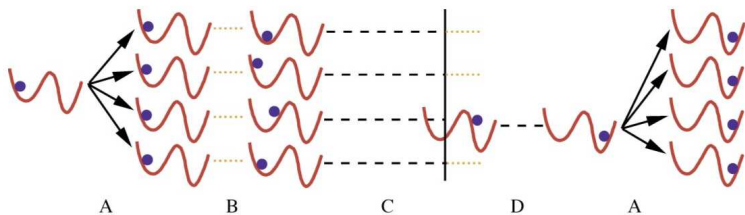
If the state is metastable, the QSD is a good intermediate between continuous-state space dynamics and jump Markov models.

We are working on generalizations:

- Broader geometric setting
- Langevin dynamics
- Non-reversible dynamics

The mathematical analysis gives the proper geometric setting under which the kinetic Monte Carlo model can be built and the Eyring-Kramers formulas can be used to parameterize it.

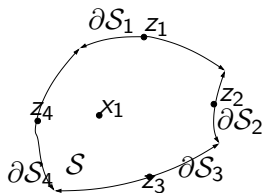
2- From theory to algorithms



A.F. Voter, Annu. Rev. Mater. Res., vol. 32, 2002.

How to sample efficiently the exit event?

If the process remains sufficiently long in a state, the exit event can be modeled by one jump of a Markov state model. This can be used to simulate efficiently the exit event: accelerated dynamics *à la* A.F. Voter.



Two steps:

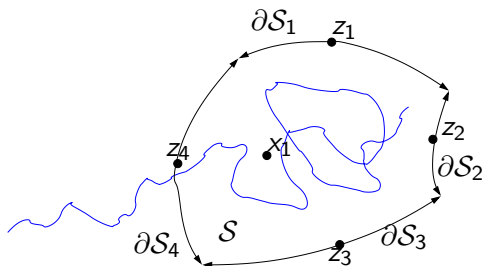
- Estimate the **decorrelation time**, namely the time to reach the QSD
- Use the underlying jump Markov process to efficiently sample the exit event

Decorrelation time

How long should we wait in practice so that $\mathcal{L}(\mathbf{X}_t | \tau_S > t)$ is close to the QSD ν ?

- Theoretically: exponential decay

$$\|\mathcal{L}(\mathbf{X}_t | \tau_S > t) - \nu\|_{TV} \leq C(\mathcal{L}(\mathbf{X}_0)) \exp(-(\lambda_2 - \lambda_1)t);$$
- Numerically: simulate $\mathcal{L}(\mathbf{X}_t | \tau_S > t)$ via interacting particle system (Fleming-Viot particle system), and test stationarity to estimate the convergence time to the QSD (Gelman-Rubin convergence diagnostic).



The Fleming-Viot particle process

Start N processes i.i.d. from μ_0 , and iterate the following steps:

1. Integrate (in parallel) N realizations ($k = 1, \dots, N$)

$$d\mathbf{X}_t^k = -\nabla V(\mathbf{X}_t^k) dt + \sqrt{2\beta^{-1}} d\mathbf{W}_t^k$$

until one of them, say \mathbf{X}_t^1 , exits;

2. Kill the process that exits;
3. With uniform probability $1/(N-1)$, randomly choose one of the survivors, $\mathbf{X}_t^2, \dots, \mathbf{X}_t^N$, say \mathbf{X}_t^2 ;
4. Branch \mathbf{X}_t^2 , with one copy persisting as \mathbf{X}_t^2 , and the other becoming the new \mathbf{X}_t^1 .

It is known that the empirical distribution [Villemonais]

$$\mu_{t,N} \equiv \frac{1}{N} \sum_{k=1}^N \delta_{\mathbf{X}_t^k}$$

satisfies:

$$\lim_{N \rightarrow \infty} \mu_{t,N} = \mathcal{L}(\mathbf{X}_t | t < \tau_S).$$

Accelerated dynamics

Once the QSD has been reached, there are three ideas to efficiently sample $(\tau_S, \mathbf{X}_{\tau_S})$:

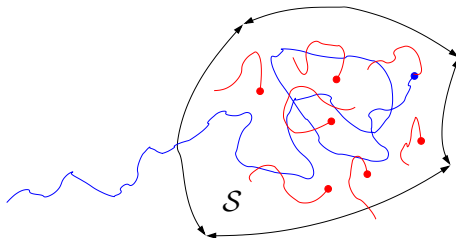
- use parallel architectures to accelerate the sampling: parallel replica, parsplicing
- raise the minimum of the potential inside the state \mathcal{S} (but not on $\partial\mathcal{S}$): hyperdynamics
- raise the temperature: temperature accelerated dynamics

The Parallel Replica Algorithm

Perform many independent exit events **in parallel** [Voter, 1998]

Two steps:

- Distribute N independent initial conditions in \mathcal{S} according to the QSD ν ;
- Evolve N replicas from these initial conditions, consider **the first exiting replica**, and multiply the first exit time by the number of replicas.



The Parallel Replica Algorithm

Why is it consistent?

- Exit time is independent of exit point so that

$$\mathbf{X}_{\tau_S^{l_0}}^{l_0} \stackrel{\mathcal{L}}{=} \mathbf{X}_{\tau_S^1}^1,$$

where $l_0 = \arg \min_i (\tau_S^i)$;

- Exit times are i.i.d. exponentially distributed so that, for all N ,

$$N \min(\tau_S^1, \dots, \tau_S^N) \stackrel{\mathcal{L}}{=} \tau_S^1.$$

Remark: For this algorithm, one just needs two properties: τ_S is exponentially distributed, and independent of the exit point \mathbf{X}_{τ_S} . The Eyring-Kramers formulas are not used.

The original Parallel Replica Algorithm

The **original** parallel replica algorithm is in three steps:

- **Decorrelation step** with a fixed deterministic decorrelation time τ_{corr}
- **Dephasing step** to get N initial conditions i.i.d. with law the QSD
- **Parallel step**

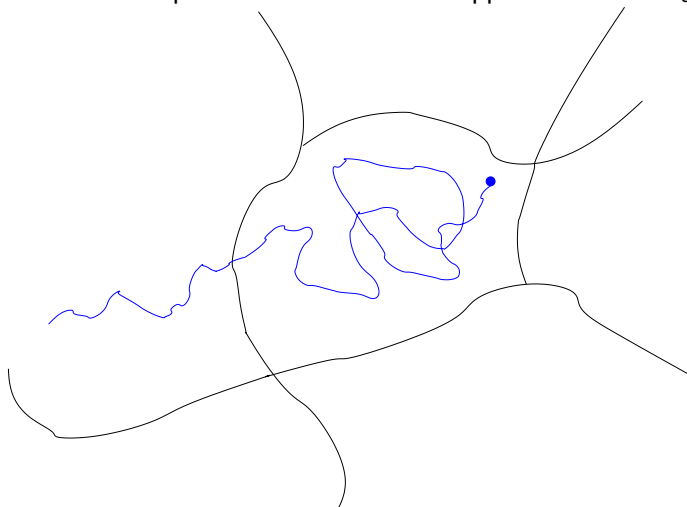
The original Parallel Replica Algorithm

Decorrelation step: run the dynamics on a reference walker...



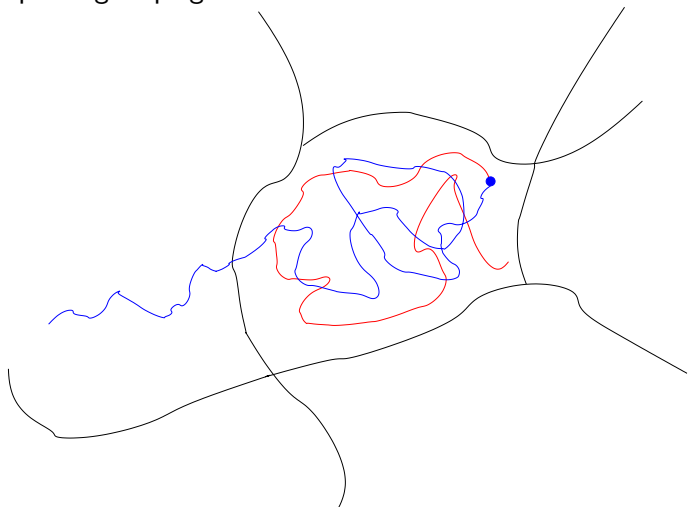
The original Parallel Replica Algorithm

Decorrelation step: ... until it remains trapped for a time τ_{corr} .



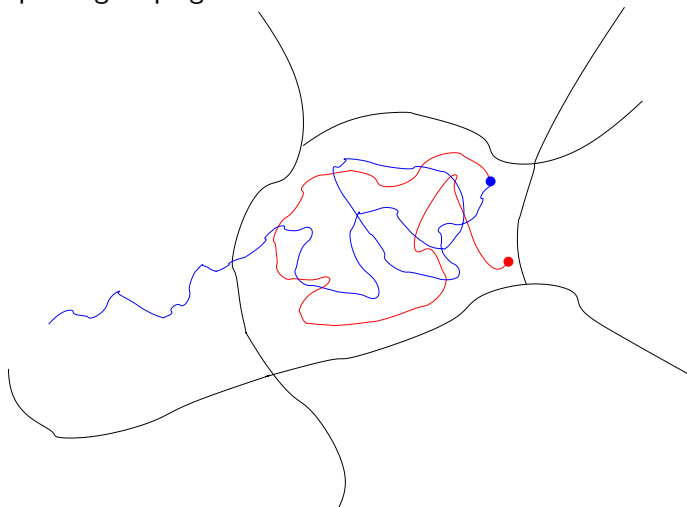
The original Parallel Replica Algorithm

Dephasing step: generate new initial conditions in the state.



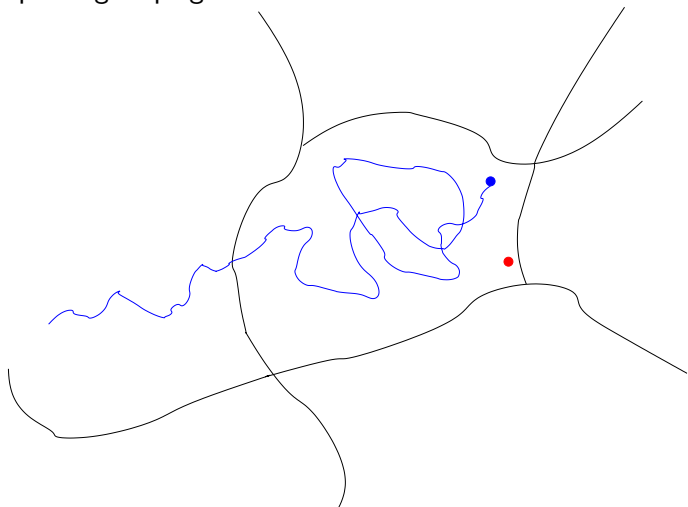
The original Parallel Replica Algorithm

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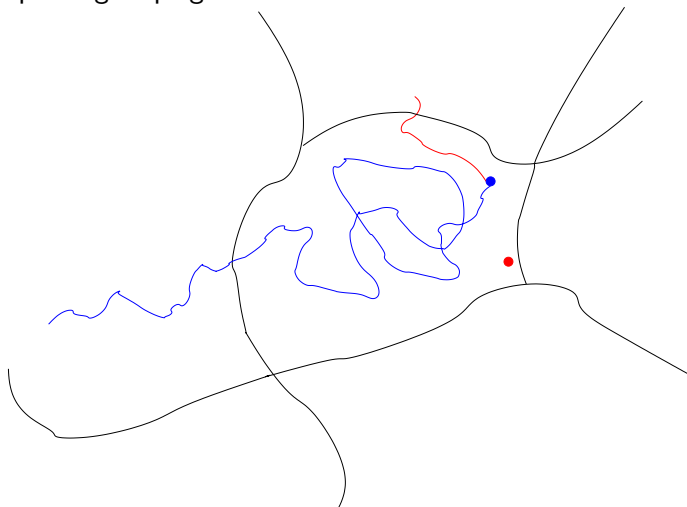
The original Parallel Replica Algorithm

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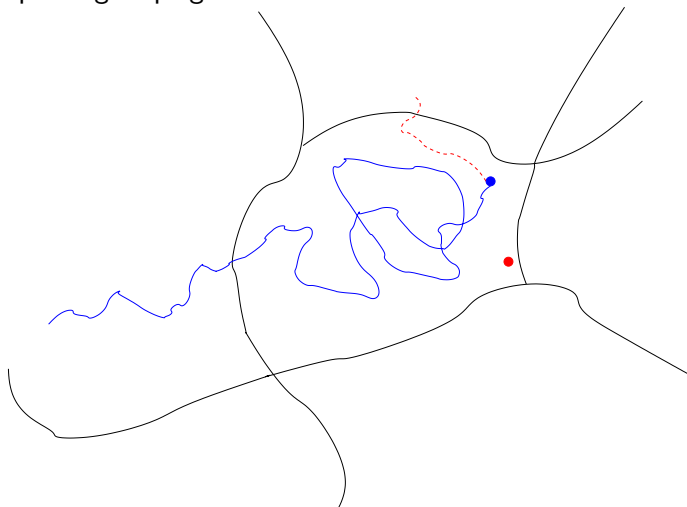
The original Parallel Replica Algorithm

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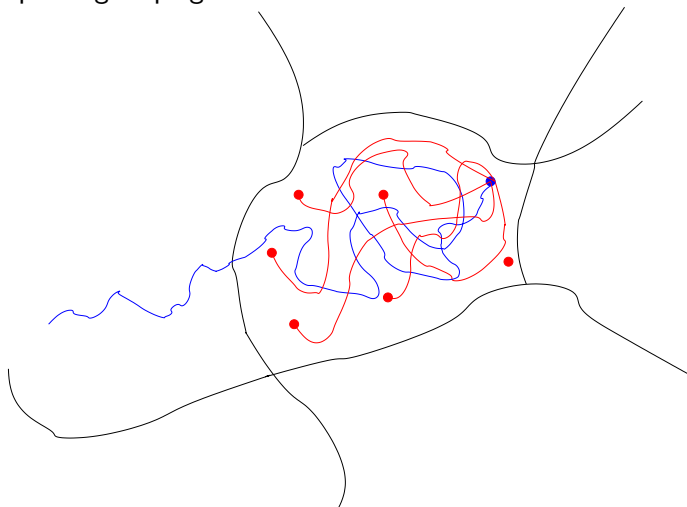
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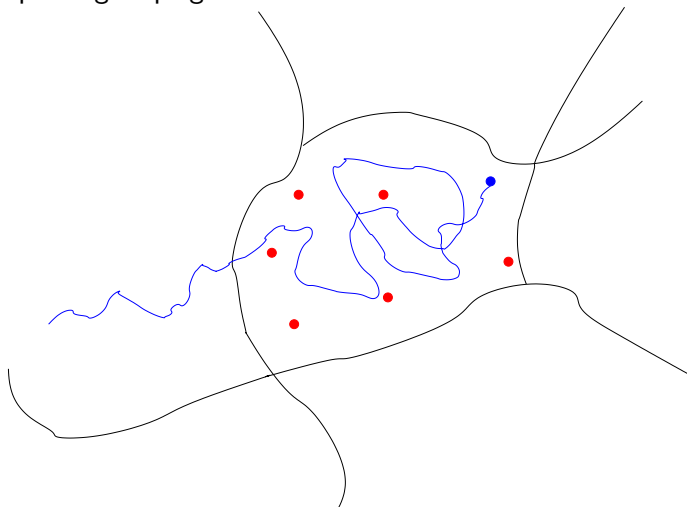
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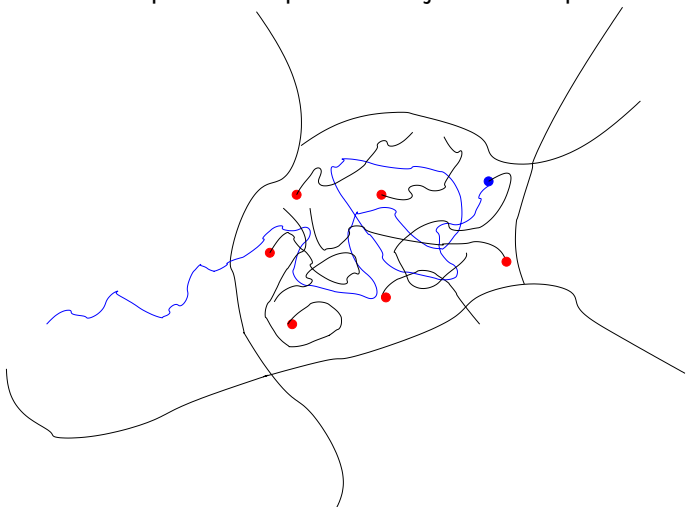
The original Parallel Replica Algorithm

Dephasing step: generate new initial conditions in the state.



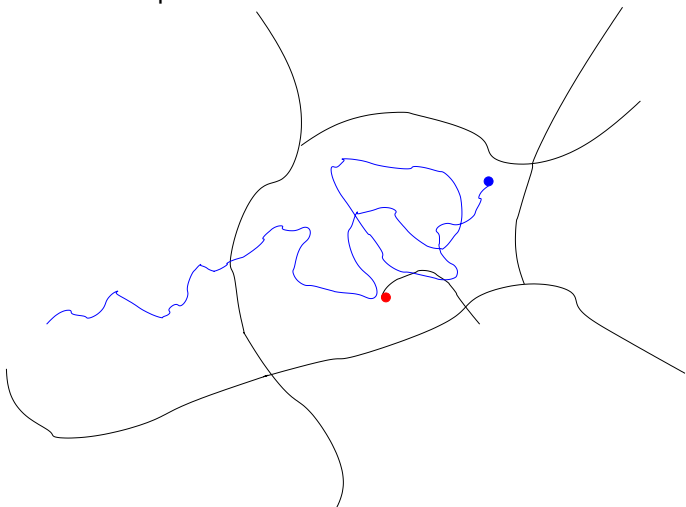
The original Parallel Replica Algorithm

Parallel step: run independent trajectories in parallel...



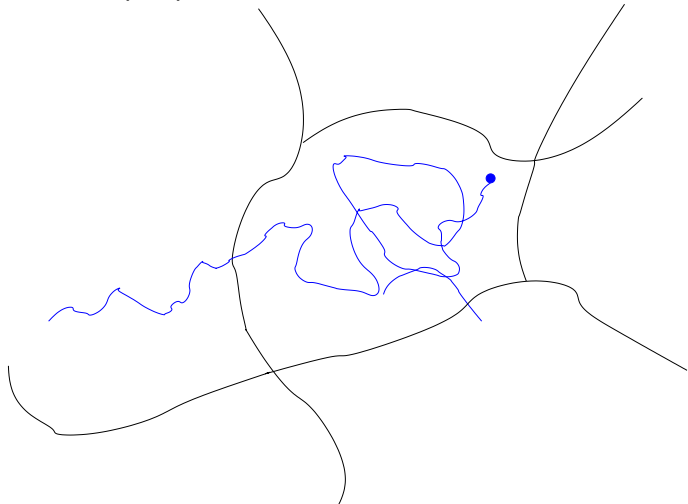
The original Parallel Replica Algorithm

Parallel step: ... and detect the first transition event.



The original Parallel Replica Algorithm

Parallel step: update the time clock: $T_{simu} = T_{simu} + NT$.



The original Parallel Replica Algorithm

A new decorrelation step starts...



The generalized Parallel Replica algorithm

[Binder, Hédin, TL, Simpson, 2015]

1. Run a reference walker, using standard MD.
2. Each time the reference walker enters a state, start a Fleming-Viot particle process (with N replicas simulated in parallel) with initial condition the entering point.
3. If the reference walker exits before the Fleming Viot particle process reaches stationarity go back to 1. Else go to the parallel step.
4. Parallel step: Starting from the end points of the Fleming-Viot particle process (approximately i.i.d. with law the QSD), run independent MD and consider the first exit event. Multiply the first exit time by N and go back to 1, using the first exit point as initial condition.

The time at which the Fleming-Viot particle process becomes stationary is determined using the Gelman-Rubin statistical test.

The generalized Parallel Replica algorithm

- The algorithm does not require a partition of the state space but only an ensemble of states.
- The time to reach the QSD is estimated each time the process enters a new state (it depends on the state and on the initial condition within the state).

Numerical results

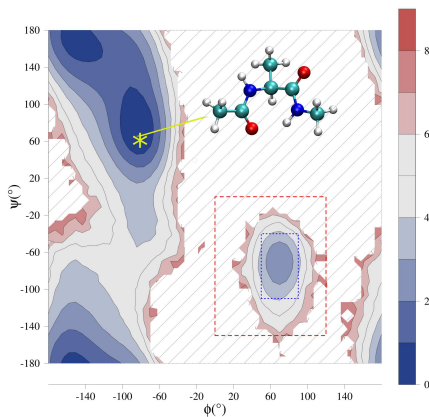
We tested the generalized Parallel Replica algorithm applied to biological systems (with Florent Hédin):

- Conformational equilibrium of the alanine dipeptide
- Dissociation of the FKBP-DMSO protein-ligand system

Main differences with materials science: definition of the states using collective variables, the states do not define a partition, much more rugged landscapes.

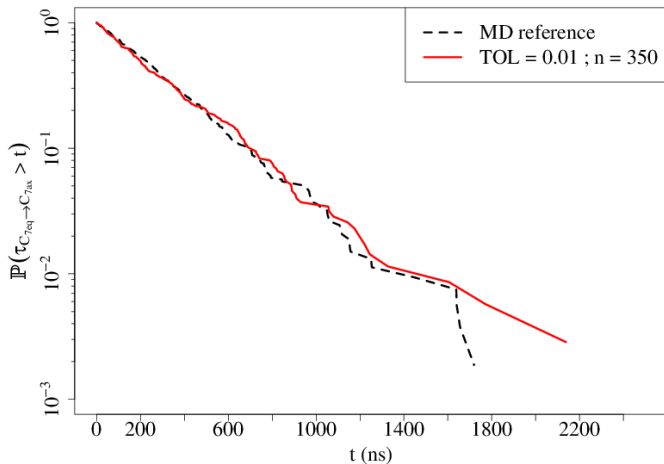
Current implementation within OpenMM, see <https://gitlab.inria.fr/parallel-replica>

Alanine dipeptide (1/5)



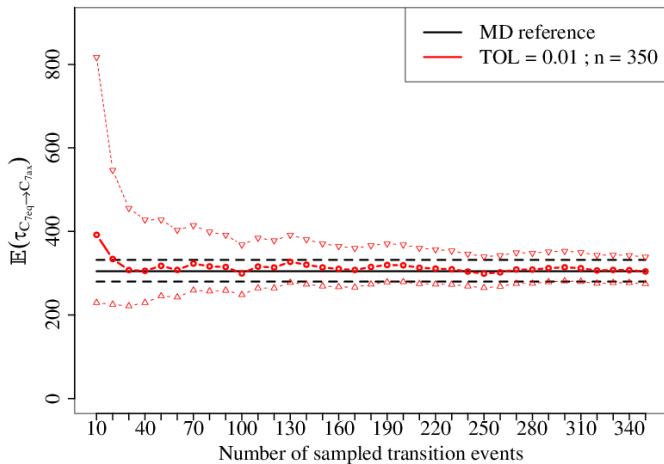
Definition of ParRep domains based on a free energy surface: we study the transition time from C_{7eq} (outside the red rectangle) to C_{7ax} (inside the red rectangle).

Alanine dipeptide (2/5)



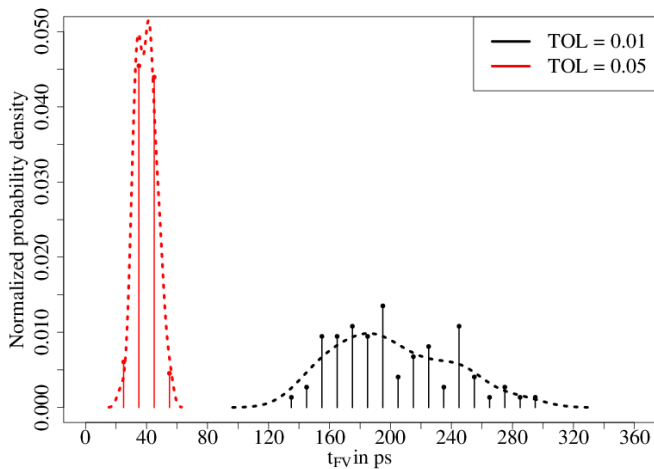
Cumulative distribution function of the transition time.

Alanine dipeptide (3/5)



Convergence of the mean transition time.

Alanine dipeptide (4/5)



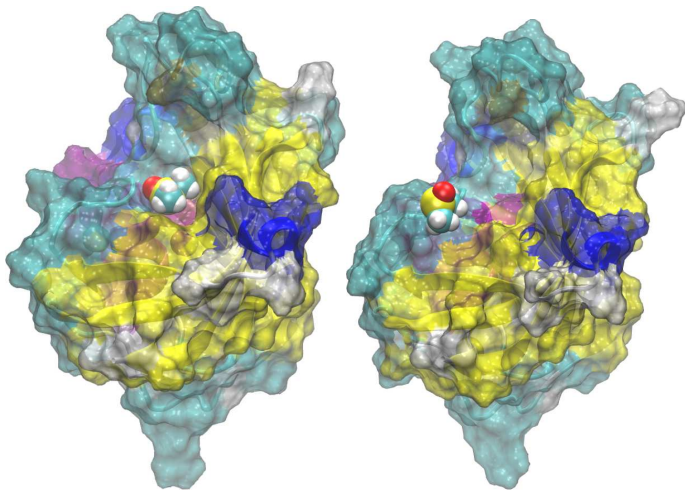
Distribution of the correlation times computed by FV.

Alanine dipeptide (5/5)

tol	WT(s)	t_{sim} (ns)	Speed(ns/day)	Eff. speedup	(Eff./Max)
0.01	6015	10008	143752	156	70%
0.025	5239	10103	166609	181	80%
0.05	4973	10032	174296	189	84%

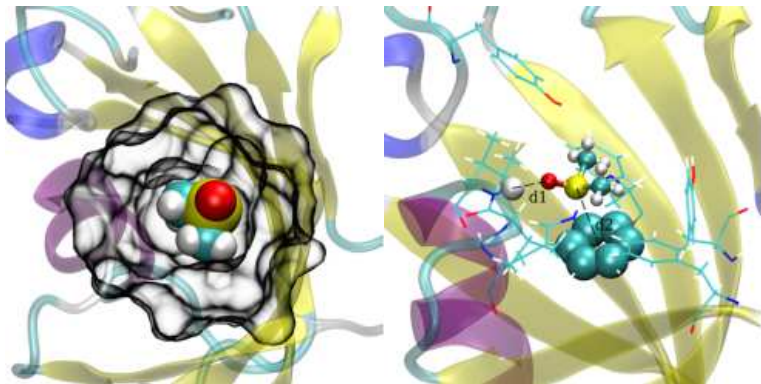
Effective speed-up as a function of the tolerance, for $N = 224$ replicas run in parallel (speed of a reference Langevin dynamics is 921 ns/day).

FKBP-DMSO (1/4)



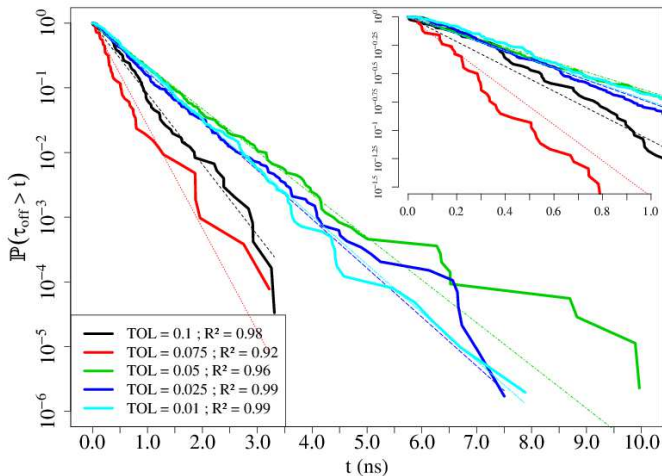
FKBP-DMSO complex,
corresponding to the RCSB-PDB entry "1D7H".

FKBP-DMSO (2/4)



DMSO in its binding cavity ; distances used to define the cavity.

FKBP-DMSO (3/4)



Cumulative distribution function of the dissociation times.

FKBP-DMSO (4/4)

TOL	WT(s)	t_{sim} (ns)	Speed (ns/day)	Eff. speedup	(Eff./Max)
0.01	85142	403.5	409.4	79.5	56.8%
0.025	79574	457.6	496.8	96.5	68.9%
0.05	84455	482.2	493.4	95.8	68.4%

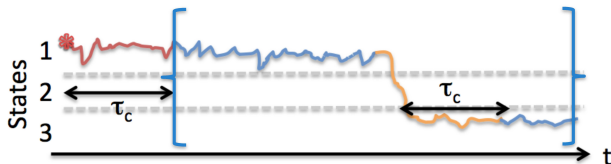
Effective speed-up as a function of the tolerance, for $N = 140$ replicas run in parallel (speed of a reference Langevin dynamics is 5.15 ns/day).

The Parallel Trajectory Splicing algorithm

Precompute the exit events [Perez, Cubuk, Waterland, Kaxiras, Voter, 2015]

Algorithm:

- Simulate in parallel short trajectories which start from the QSD in a state, and end at the QSD in a state.
- Glue together these short trajectories to build the full dynamics.

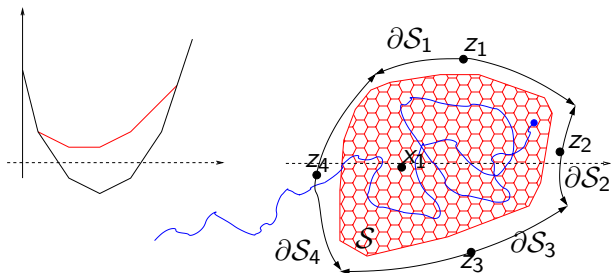


Hyperdynamics (1/2)

Raise the potential in \mathcal{S} to reduce the exit time [Voter, 1997]

Two steps:

- Equilibrate on the **biased potential** $V + \delta V$;
- Wait for an exit and multiply the exit time $\tau_S^{\delta V}$ by the boost factor $B = \frac{1}{\tau_S^{\delta V}} \int_0^{\tau_S^{\delta V}} \exp(\beta \delta V(\mathbf{X}_t)) dt$.



Hyperdynamics (2/2)

Why is it consistent ?

Assumptions on δV : (i) $\delta V = 0$ on ∂S and (ii) δV is sufficiently small so that the Theorem above applies.

Recall the formula for the exit rates:

$$k_i = \tilde{\nu}_i^{OL} e^{-\beta[V(z_i) - V(x_1)]} (1 + O(\beta^{-1}))$$

where $\tilde{\nu}_i^{OL} = \sqrt{\frac{\beta}{2\pi}} \partial_n V(z_i) \frac{\sqrt{\det(\nabla^2 V)(x_1)}}{\sqrt{\det(\nabla^2 V|_{\partial S})(z_i)}}$.

One easily check that $k_i / \sum_{j=1}^J k_j$ is independent of δV and

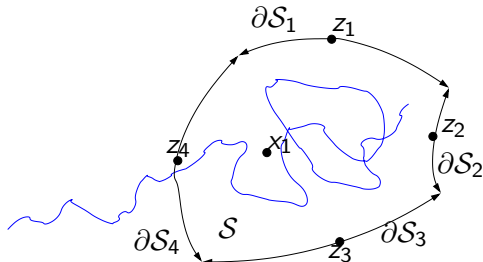
$$\begin{aligned} \frac{\sum_{j=1}^J k_j(V + \delta V)}{\sum_{j=1}^J k_j(V)} &= \sqrt{\frac{\det(\nabla^2(V + \delta V))(x_1)}{\det(\nabla^2(V))(x_1)}} e^{\beta\delta V(x_1)} (1 + O(\beta^{-1})) \\ &= \frac{\int_S \exp(-\beta V)}{\int_S \exp(-\beta(V + \delta V))} (1 + O(\beta^{-1})) \simeq B \end{aligned}$$

Temperature Accelerated Dynamics (1/2)

Increase the temperature to reduce the exit time [Sorensen, Voter, 2000]

Algorithm:

- Observe the exit events from \mathcal{S} at high temperature ;
- Extrapolate the high temperature exit events to low temperature exit events.



Temperature Accelerated Dynamics (2/2)

Recall that, starting from the QSD, the exit event from a given state \mathcal{S} can exactly be modelled using a kinetic Monte Carlo model with rates

$$k_i = \tilde{\nu}_i^{OL} e^{-\beta[V(z_i) - V(x_1)]} (1 + O(\beta^{-1}))$$

where $\tilde{\nu}_i^{OL} = \sqrt{\frac{\beta}{2\pi}} \partial_n V(z_i) \frac{\sqrt{\det(\nabla^2 V)(x_1)}}{\sqrt{\det(\nabla^2 V|_{\partial\mathcal{S}})(z_i)}}$.

Thus,

$$\frac{k_i^{lo}}{k_i^{hi}} \simeq \sqrt{\frac{\beta^{lo}}{\beta^{hi}}} \exp(-(\beta^{lo} - \beta^{hi})(V(z_i) - V(x_1))).$$

Algorithm: observe exit events at high temperature, extrapolate the rates to low temperature, **stop when the extrapolated event will not modify anymore the low temperature exit event.**

Remark: TAD can be seen as a smart saddle point search method.

Generalizations and perspectives

- The parallel replica is a very versatile algorithm: it applies e.g. to non reversible dynamics, discrete-in-time dynamics, continuous-time Markov Chain [Aristoff, Plechac, Wang]. It does not require estimates of the exit rates.
- Hyper and TAD are more efficient, but require the temperature to be sufficiently small so that estimates of the rates by the Eyring-Kramers formulas hold true.

All these techniques require “good” metastable states:
exit time $>$ convergence time to the QSD.

Conclusion

There are mathematical characterizations of good coarse-graining representations (spectral gaps, convergence times vs exit times).

Could we use those characterizations together with advanced learning techniques (auto-encoder, sparse methods) to get better coarse-grained descriptions?

- Identify slow variables
- Sparse representation of the committor function
- Identify metastable states

References

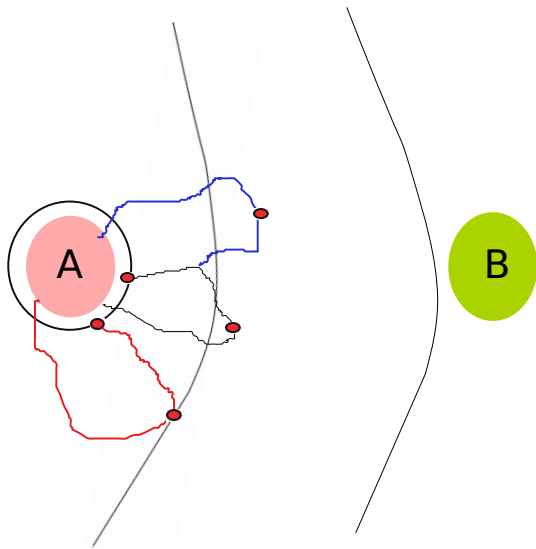
Some papers I mentioned:

- A. Binder, TL and G. Simpson, *A Generalized Parallel Replica Dynamics*, Journal of Computational Physics, 284, 2015.
- G. Di Gesù, TL, D. Le Peutrec and B. Nectoux, *Jump Markov models and transition state theory: the Quasi-Stationary Distribution approach*, Faraday Discussion, 195, 2016.
- G. Di Gesù, TL, D. Le Peutrec and B. Nectoux, *Sharp asymptotics of the first exit point density*, Annals of PDE, 5, 2019.
- F. Hédin and TL, *gen.parRep: a first implementation of the Generalized Parallel Replica dynamics for the long time simulation of metastable biochemical systems*, Computer Physics Communications, 239, 2019.
- C. Le Bris, TL, M. Luskin and D. Perez, *A mathematical formalization of the parallel replica dynamics*, Monte Carlo Methods and Applications, 18(2), 2012.

Splitting strategies

- 1- The Adaptive Multilevel Splitting algorithm
- 2- Computing transition times with AMS

1- The Adaptive Multilevel Splitting algorithm



Multilevel splitting

General setting: Let $(\mathbf{X}_t)_{t \geq 0}$ be a Markovian dynamics, and τ_B and τ_A two associated stopping times.

Objective: efficiently compute quantities of the form $\mathbb{E}[F((\mathbf{X}_t)_{0 \leq t \leq \tau_A \wedge \tau_B}) \mathbf{1}_{\tau_B < \tau_A}]$ in the rare event setting:

$$\mathbb{P}(\tau_B < \tau_A) \ll 1.$$

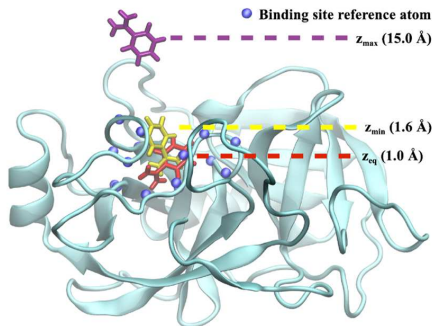
Two examples:

- Reactive trajectories: A and B are two metastable states, τ_A and τ_B are the first hitting time of A and B .
- Killed process: τ_A is a killing time, τ_B is the first hitting time of a domain B .

Motivation 1: Simulations of biological systems

Unbinding of a ligand from a protein

Trypsin with various conformational states of benzamidine



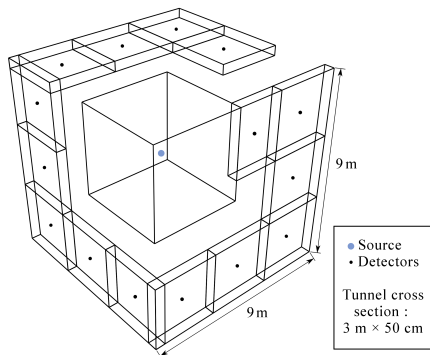
Elementary time-step for the molecular dynamics = 10^{-15} s

Dissociation time $\simeq 0.02$ s

Challenge: bridge the gap between timescales

Motivation 2: Radiation protection

Monte Carlo particle transport



Concrete tunnel with a neutron source

How to compute the neutron flux at the detector ?

Challenge: the flux is very small

Multilevel splitting: the reactive trajectory setting

We would like to sample trajectories between two given metastable states A and B . The main assumption is that we are given a smooth one dimensional function $\xi : \mathbb{R}^d \rightarrow \mathbb{R}$ which "indexes" the transition from A to B in the following sense:

$$A \subset \{x \in \mathbb{R}^d, \xi(x) < z_{\min}\} \text{ and } B \subset \{x \in \mathbb{R}^d, \xi(x) > z_{\max}\},$$

where $z_{\min} < z_{\max}$, and $\Sigma_{z_{\min}}$ (resp. $\Sigma_{z_{\max}}$) is "close" to ∂A (resp. ∂B).

Example: $\xi(x) = \|x - x_A\|$ where x_A is a reference configuration in A .

We are interested in the event $\{\tau_A < \tau_B\}$, starting from an initial condition with support in $\{x \in \mathbb{R}^d, \xi(x) < z_{\min}\}$, where

$$\tau_A = \inf\{t > 0, \mathbf{X}_t \in A\}, \quad \tau_B = \inf\{t > 0, \mathbf{X}_t \in B\}.$$

Multilevel splitting

Objective: Simulate efficiently trajectories which reach B before A and estimate $\mathbb{P}(\tau_B < \tau_A)$. This then gives dynamical information: reactive trajectories from A to B , transition times from A to B , ...

We present a **multilevel splitting approach** [Kahn, Harris, 1951] [Rosenbluth, 1955] to discard failed trajectories and branch trajectories approaching the rare set. We focus on an adaptive variant [C erou, Guyader, 2007] [C erou, Guyader, TL, Pommier, 2010]: the **Adaptive Multilevel Splitting** (AMS) algorithm.

Remark: The algorithm can be seen as a kind of adaptive Forward Flux Sampling [Allen, Valeriani, Ten Wolde, 2009]. It is also related to the Interface Sampling Method [Bolhuis, van Erp, Moroni 2003] and the Milestoning method [Elber, Faradjian 2004]. See the review paper [Bolhuis, Dellago, 2009]

Reactive trajectory

A **reactive trajectory** between two metastable sets A and B is a piece of equilibrium trajectory that leaves A and goes to B without going back to A in the meantime [Hummer,2004] [Metzner, Schütte, Vanden-Eijnden, 2006].



Difficulty: A trajectory leaving A is more likely to go back to A than to reach B .

Splitting algorithm: basic idea

The idea of splitting algorithms (FFS, RESTART, ...) is to write the rare event

$$\{\tau_B < \tau_A\}$$

as a sequence of nested events: for $z_{\min} = z_1 < \dots < z_Q = z_{\max}$,

$$\{\tau_{z_1} < \tau_A\} \supset \{\tau_{z_2} < \tau_A\} \supset \dots \supset \{\tau_{z_{\max}} < \tau_A\} \supset \{\tau_B < \tau_A\}$$

where $\tau_z = \inf\{t > 0, \xi(\mathbf{X}_t) > z\}$ and to simulate the successive *conditional events*: for $q = 1, \dots, Q - 1$,

$$\{\tau_{z_{q+1}} < \tau_A\} \text{ knowing that } \{\tau_{z_q} < \tau_A\}.$$

It is then easy to build an unbiased estimator of

$$\mathbb{P}(\tau_B < \tau_A) = \mathbb{P}(\tau_{z_1} < \tau_A) \mathbb{P}(\tau_{z_2} < \tau_A | \tau_{z_1} < \tau_A) \dots \mathbb{P}(\tau_B < \tau_A | \tau_{z_{\max}} < \tau_A)$$

Splitting algorithm: adaptive level computation

Problem: How to choose the intermediate levels $(z_q)_{q \geq 1}$?

In an ideal setting, for a given number of intermediate levels, the optimum in terms of variance is attained if

$$\forall q \geq 1, \mathbb{P}(\tau_{z_q} < \tau_A | \tau_{z_{q-1}} < \tau_A) = \mathbb{P}(\tau_{z_2} < \tau_A | \tau_{z_1} < \tau_A).$$

This naturally leads to an adaptive version (AMS, nested sampling) where the levels are determined by using *empirical quantiles*: Fix $k < n$; at iteration $q \geq 1$, given n trajectories $(\mathbf{X}_{t \wedge \tau_A}^\ell)_{t > 0, \ell = 1, \dots, n}$ in the event $\{\tau_{z_{q-1}} < \tau_A\}$, choose z_q so that

$$\mathbb{P}(\tau_{z_q} < \tau_A | \tau_{z_{q-1}} < \tau_A) \simeq \left(1 - \frac{k}{n}\right).$$

The level z_q is **the k -th order statistics of $\sup_{t \geq 0} \xi(\mathbf{X}_{t \wedge \tau_A}^\ell)$** :

$$\sup_{t \geq 0} \xi(\mathbf{X}_{t \wedge \tau_A}^{(1)}) < \dots < \sup_{t \geq 0} \xi(\mathbf{X}_{t \wedge \tau_A}^{(k)}) =: z_q < \dots < \sup_{t \geq 0} \xi(\mathbf{X}_{t \wedge \tau_A}^{(n)}).$$

AMS: estimator of the rare event probability (1/2)

Let Q_{iter} be the number of iterations to reach the level z_{max} :

$$Q_{\text{iter}} = \min\{q \geq 0, z_q > z_{\text{max}}\}$$

(where z_0 is the k -th order statistics of the n initial trajectories). Then, one obtains the estimator:

$$\left(1 - \frac{k}{n}\right)^{Q_{\text{iter}}} \simeq \mathbb{P}(\tau_{z_{\text{max}}} < \tau_A).$$

AMS: estimator of the rare event probability (2/2)

At iteration Q_{iter} , one has an ensemble of n trajectories such that $\tau_{z_{\max}} < \tau_A$. Thus

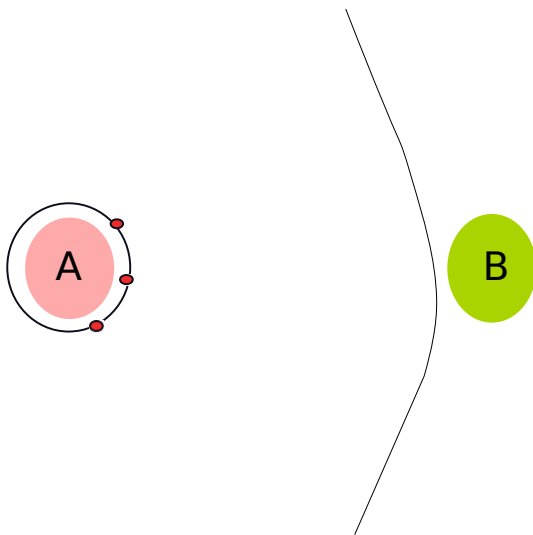
$$\hat{p}_{\text{corr}} := \frac{1}{n} \sum_{\ell=1}^n 1_{\{\tau_B(\mathbf{x}^{\ell, Q_{\text{iter}}}) < \tau_A(\mathbf{x}^{\ell, Q_{\text{iter}}})\}} \simeq \mathbb{P}(\tau_B < \tau_A | \tau_{z_{\max}} < \tau_A).$$

\hat{p}_{corr} is the number of trajectories reaching B before A at the last iteration Q_{iter} .

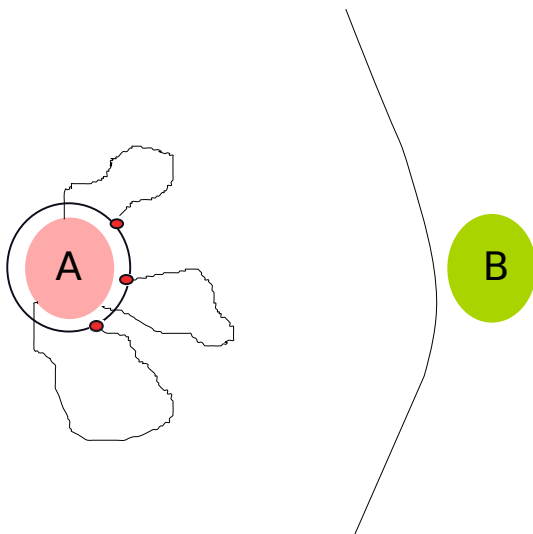
Therefore, an estimator of $\mathbb{P}(\tau_B < \tau_A)$ is

$$\left(1 - \frac{k}{n}\right)^{Q_{\text{iter}}} \hat{p}_{\text{corr}}.$$

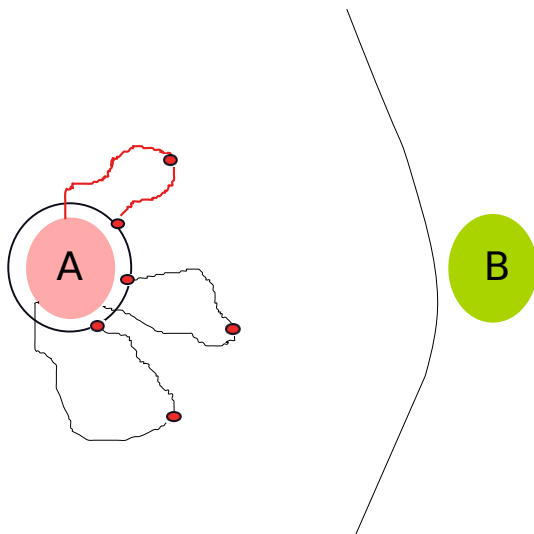
AMS Algorithm



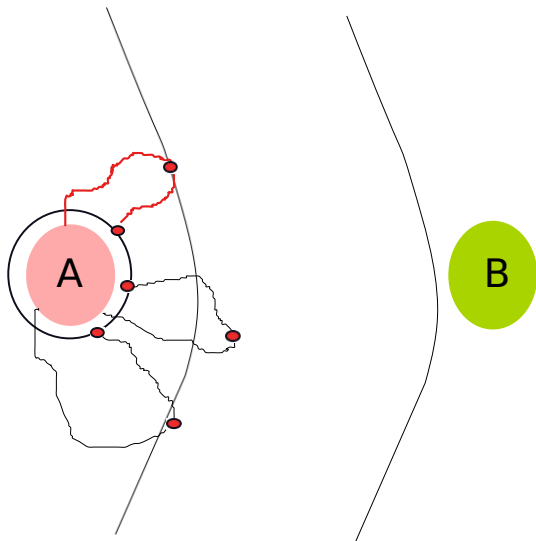
AMS Algorithm



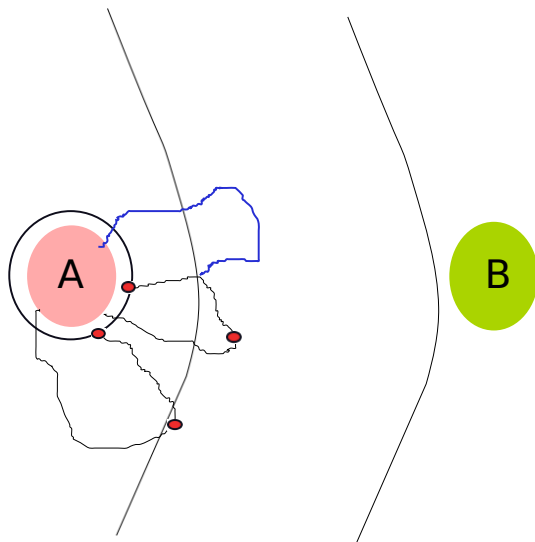
AMS Algorithm



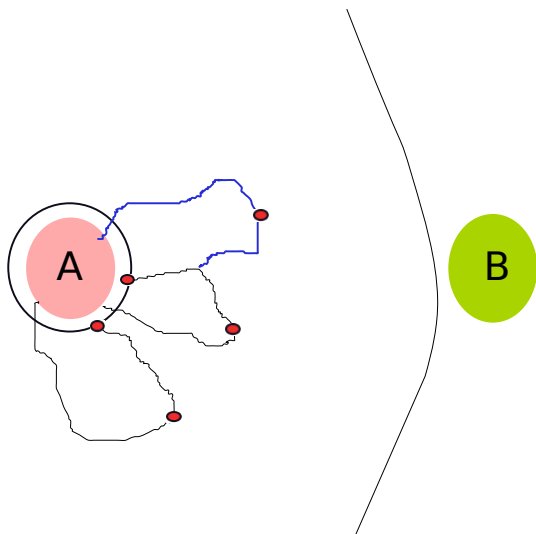
AMS Algorithm



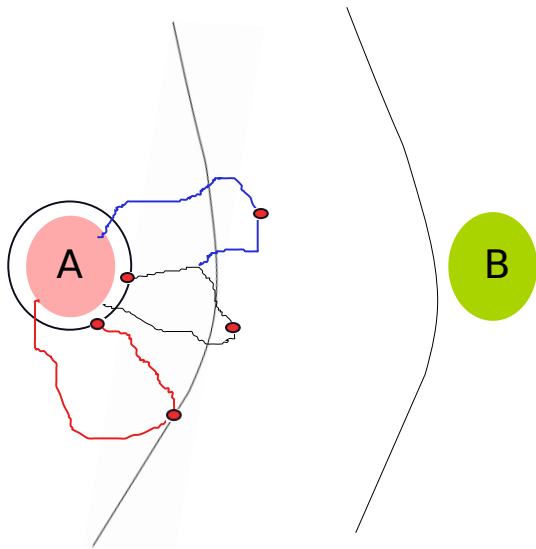
AMS Algorithm



AMS Algorithm



AMS Algorithm



AMS Algorithm: the case of Markov chains

In practice, the dynamics are *discrete in time* and thus, it may happen that more than k trajectories are such that

$$\sup_{t \geq 0} \xi(\mathbf{X}_{t \wedge \tau_A}^\ell) \leq \sup_{t \geq 0} \xi(\mathbf{X}_{t \wedge \tau_A}^{(k)}) =: z_q$$

In this case, **all the trajectories with maximum level smaller or equal than z_q should be discarded.**

The actual estimator of $\mathbb{P}(\tau_B < \tau_A)$ thus reads:

$$\hat{p} = \left(1 - \frac{K_1}{n}\right) \dots \left(1 - \frac{K_{Q_{\text{iter}}}}{n}\right) \hat{p}_{\text{CORR}}$$

instead of $\left(1 - \frac{k}{n}\right)^{Q_{\text{iter}}} \hat{p}_{\text{CORR}}$, where $K_q \geq k$ is the effective number of discarded trajectories at iteration q .

AMS Algorithm: unbiasedness

Theorem [C.-E. Bréhier, M. Gazeau, L. Goudenège, TL, M. Rousset, 2016]: For any choice of ξ , n and k ,

$$\mathbb{E}(\hat{\rho}) = \mathbb{P}(\tau_B < \tau_A).$$

The proof is based on Doob's stopping theorem on a martingale built using filtrations indexed by the level sets of ξ . Actually, this result is proved for general path observables and in a much more general setting.

Practical counterparts:

- The algorithm is easy to parallelize.
- One can compare the results obtained with different reaction coordinates ξ to gain confidence in the results.

Computing transition times

To use the algorithm to compute transition times, we split a transition path from A to B into: excursions from ∂A to $\Sigma_{z_{\min}}$ and then back to ∂A , and finally an excursion from ∂A to $\Sigma_{z_{\min}}$ and then to B . Assuming that A is metastable ($p \ll 1$), it can be shown that the equilibrium mean transition time can be approximated by (see the second part of this talk):

$$\left(\frac{1}{p} - 1\right) \Delta_{\text{Loop}} + \Delta_{\text{React}}$$

where:

- p is the probability, starting from $\Sigma_{z_{\min}}$ “at equilibrium”, to go to B rather than A (approximated by \hat{p});
- Δ_{Loop} is the mean time for an excursion from ∂A to $\Sigma_{z_{\min}}$ and then back to ∂A (approximated by brute force);
- Δ_{React} is the mean time for an excursion from ∂A to $\Sigma_{z_{\min}}$ and then to B (approximated by the AMS algorithm).

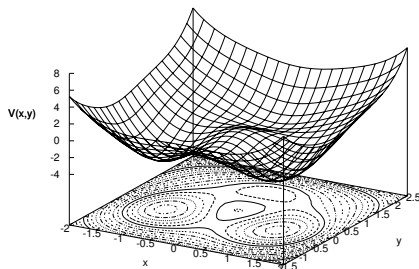
Numerical results: a 2D example

Time-discretization of the overdamped Langevin dynamics:

$$d\mathbf{X}_t = -\nabla V(\mathbf{X}_t) dt + \sqrt{2\beta^{-1}} d\mathbf{W}_t$$

with a deterministic initial condition $\mathbf{X}_0 = \mathbf{x}_0$ and the 2D potential

[Park, Sener, Lu, Schulten, 2003] [Metzner, Schütte and Vanden-Eijnden, 2006]



$$V(x, y) = 3e^{-x^2 - (y - \frac{1}{3})^2} - 3e^{-x^2 - (y - \frac{5}{3})^2} - 5e^{-(x-1)^2 - y^2} \\ - 5e^{-(x+1)^2 - y^2} + 0.2x^4 + 0.2\left(y - \frac{1}{3}\right)^4.$$

A 2D example

The interest of this “bi-channel” potential is that, depending on the temperature, one or the other channel is preferred to go from A (around $H_- = (-1, 0)$) to B (around $H_+ = (1, 0)$).

Three reaction coordinates: $\xi^1(x, y) = \|(x, y) - H_-\|$,
 $\xi^2(x, y) = C - \|(x, y) - H_+\|$ or $\xi^3(x, y) = x$.

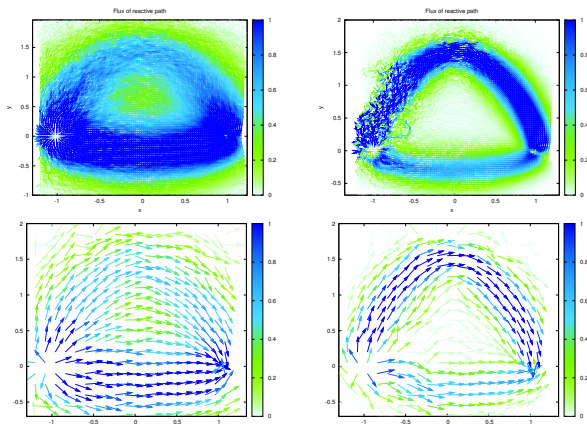
We plot as a function of the number N of independent realizations of AMS, the empirical average

$$\bar{p}_N = \frac{1}{N} \sum_{m=1}^N \hat{p}_m$$

together with the associated empirical confidence interval: $[\bar{p}_N - \delta_N/2, \bar{p}_N + \delta_N/2]$ where

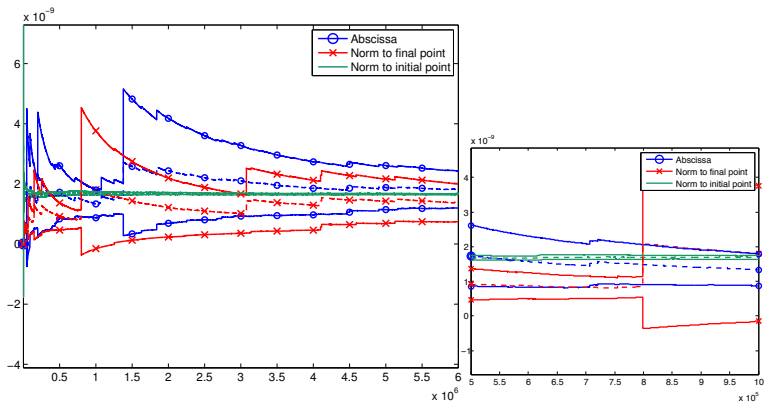
$$\delta_N = 2 \frac{1.96}{\sqrt{N}} \sqrt{\frac{1}{N} \sum_{m=1}^N (\hat{p}_m)^2 - (\bar{p}_N)^2}$$

A 2D example: flux of reactive trajectories

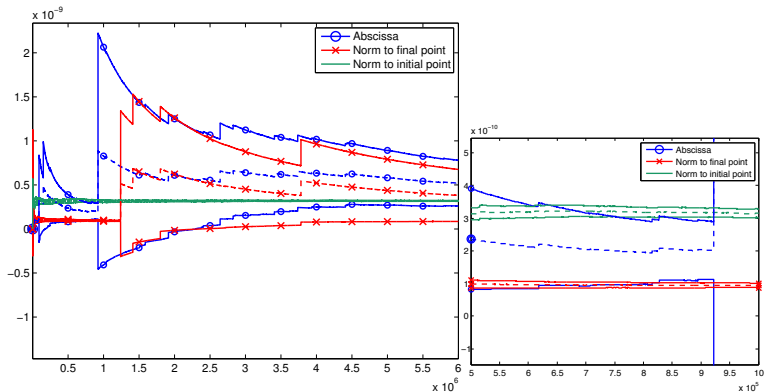


Flux of reactive trajectories, at $\beta = 1.67$ on the left, and $\beta = 6.67$ on the right.

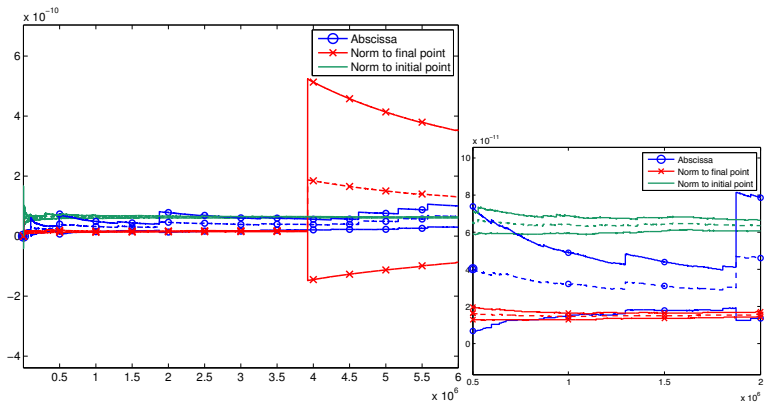
A 2D example: $k = 1$, $n = 100$, $\beta = 8.67$



A 2D example: $k = 1$, $n = 100$, $\beta = 9.33$



A 2D example: $k = 1$, $n = 100$, $\beta = 10$



A 2D example

Observations:

- When N is sufficiently large, confidence intervals overlap.
- For too small values of N , “apparent bias” is observed [Glasserman, Heidelberger, Shahabuddin, Zajic, 1998].
- Fluctuations depend a lot on ξ .

→ To gain confidence in the results, check that the estimated quantity is approximately the same for different ξ 's.

“Apparent bias” phenomenon

The apparent bias is due to the fact that [Glasserman, Heidelberger, Shahabuddin, Zajic, 1998]:

- Multiple pathways exist to go from A to B .
- Conditionally to reach Σ_z before A , the relative likelihood of each of these pathways depends a lot on z .

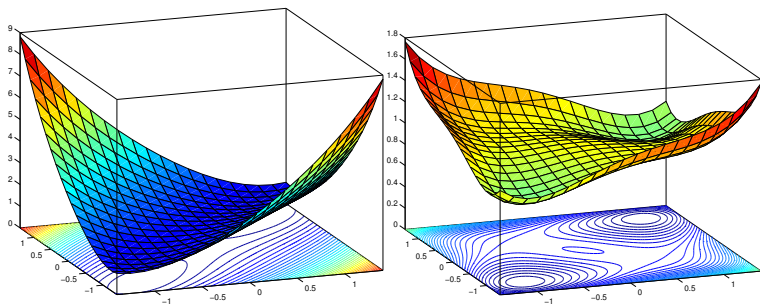
On our example, for small n , we indeed observe that (for ξ^3):

- Most of the time, all replicas at the end go through only one of the two channels (two possible scenarios).
- One of this scenario is rare.
- The values of \hat{p} associated to each of these two scenarios are very different.

This explains the large fluctuations.

“Apparent bias” phenomenon

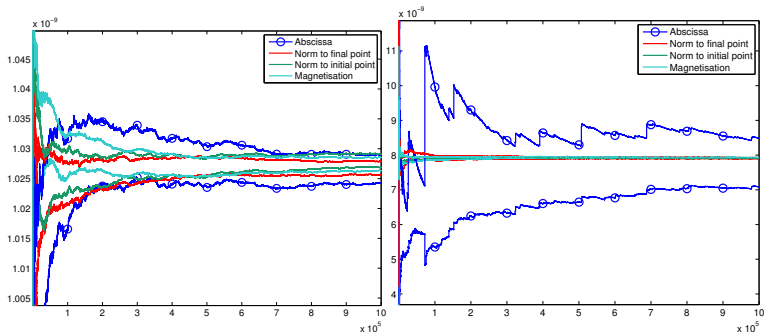
Another 2D test case:



Potential $V_\gamma(x, y)$.

Left: $\gamma = 1$ (one channel); right: $\gamma = 0.1$ (two channels).

“Apparent bias” phenomenon



Parameters: $k = 1$, $n = 100$ and $\beta = 80$.

Left: $\gamma = 1$ (one channel). Right: $\gamma = 0.1$ (two channels).

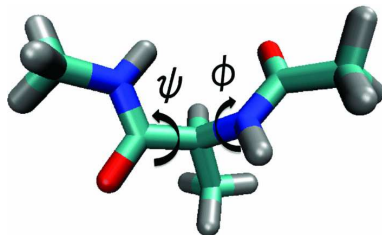
Results on larger test cases

We are currently implementing AMS in the NAMD software (collaboration with SANOFI, C. Mayne and I. Teo, PhD of L. Silva Lopes with J. Hénin).

Three test cases:

- Alanine di-peptide (test case)
- benzamidine-trypsin dissociation rate
- β -cyclodextrin (in progress)

Alanine di-peptide (1/6)

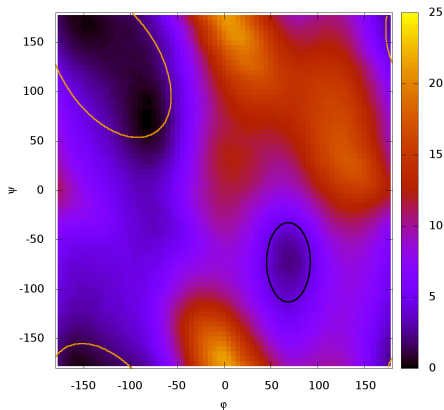


Two reaction coordinates:

- ξ_1 is a continuous piecewise affine function of φ
- $\xi_2(\varphi, \psi) = \min(d_A(\varphi, \psi), 6.4) - \min(d_B(\varphi, \psi), 3.8)$

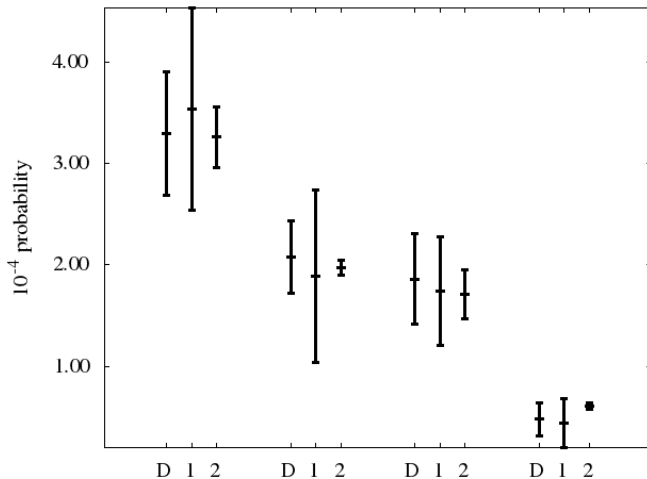
Computational setting: no solvent, force field: CHARMM27. AMS with $n = 500$ to 1000 replicas and $k = 1$.

Alanine di-peptide (2/6)



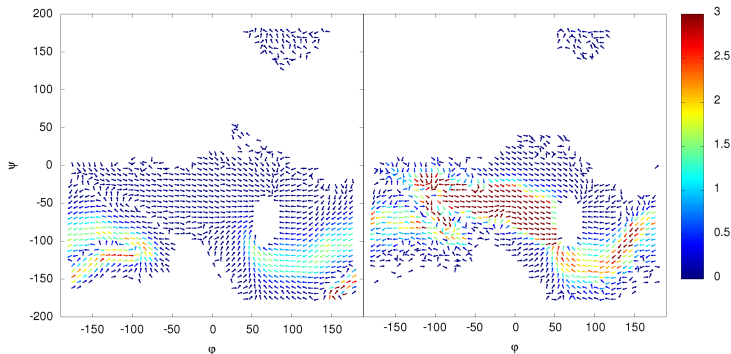
Free energy landscape and zones A (yellow) and B (black).

Alanine di-peptide (3/6)



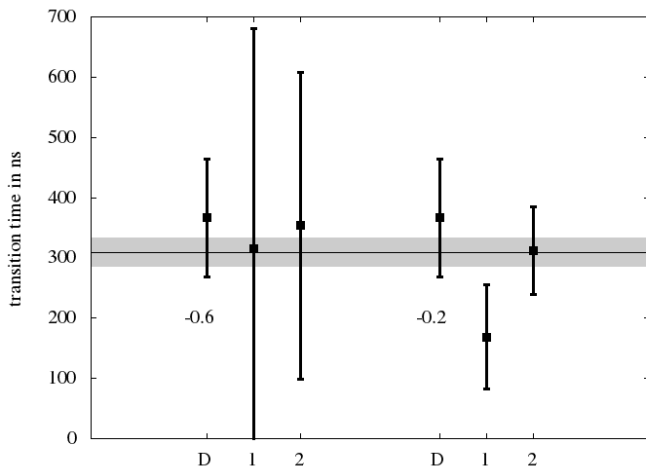
Probability estimations using different initial conditions: D=DNS,
1= ξ_1 , 2= ξ_2 .

Alanine di-peptide (4/6)



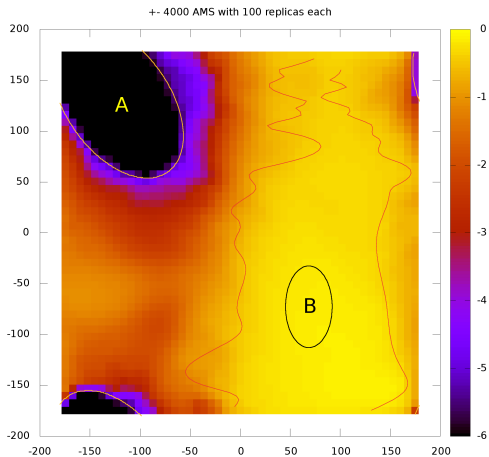
Flux of reactive trajectories, starting from two different initial conditions.

Alanine di-peptide (5/6)



Transition time obtained for two values of z_{min} : D=DNS, 1= ξ_1 , 2= ξ_2 . Reference value obtained by DNS over a 97 DNS simulations of $2\mu s$.

Alanine di-peptide (6/6)

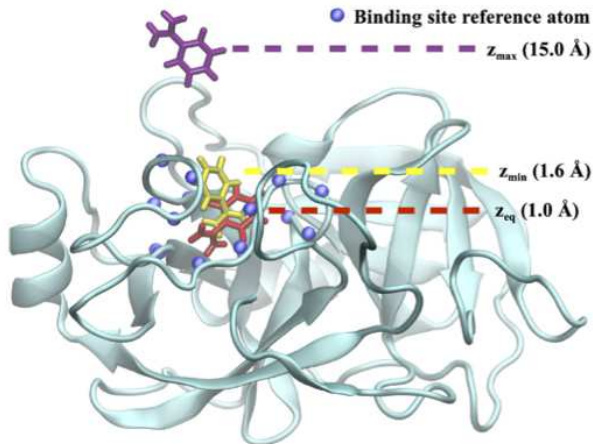


Estimate of the committor function using AMS.

Benzamidine-trypsin (1/2)

We recently used AMS to estimate the off rate of benzamidine from trypsin [I. Teo, C. Mayne, K. Schulten and TL, 2016].

Trypsin with various conformational states of benzamidine



Benzamidine-trypsin (2/2)

We obtain a dissociation rate $k_{\text{off}} = (260 \pm 240)\text{s}^{-1}$ within the same order of magnitude as the experimentally measured rate $(600 \pm 300)\text{s}^{-1}$.

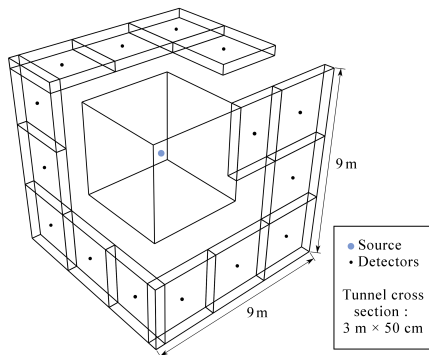
The overall simulation time taken, summed over all 1000 replicas, was $2.1\mu\text{s}$ ($2.3\mu\text{s}$ after including direct MD and steered MD simulations), which is **four orders of magnitude shorter than the estimated dissociation time of one event**.

The main practical difficulty seems to be the determination of a 'good' domain A .

Computational setting: 68 789 atoms, with 21 800 water molecules, 62 sodium ions, and 68 chloride ions. Water: TIP3P model. CHARMM36 force field, with parameters for benzamidine obtained from the CGenFF force field. NPT conditions, at 298 K and 1 atm Langevin thermostat and barostat settings, using 2 fs time steps. AMS with $n = 1000$ replicas and $k = 1$.

Another example: Radiation protection (1/2)

Monte Carlo particle transport



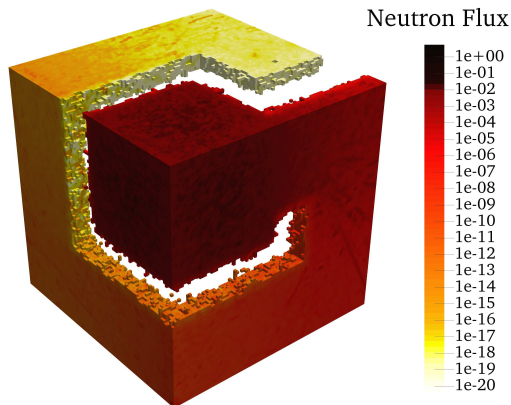
Concrete tunnel with a neutron source

How to compute the neutron flux at the detector ?

Challenge: the flux is very small

Another example: Radiation protection (2/2)

Example 2: In collaboration with CEA (Eric Dumonteil, Cheikh Diop and Henri Louvin), AMS is now implemented in the Tripoli code.



Concluding remarks on AMS (1/2)

Practical recommendations:

- A careful implementation of the splitting step leads to unbiased estimators for non-normalized quantities.
- Perform many independent realizations of AMS.
- Use ξ as a numerical parameter.

The algorithm is very versatile:

- Non-intrusivity: the MD integrator is a black box.
- Can be adapted to generate trajectories of any stopped process.
- Can be applied to both entropic and energetic barriers, to non-equilibrium systems, non-homogeneous Markov process, random fields, ...
- Algorithmic variants: other resampling procedure, additional selection, ...

Concluding remarks on AMS (2/2)

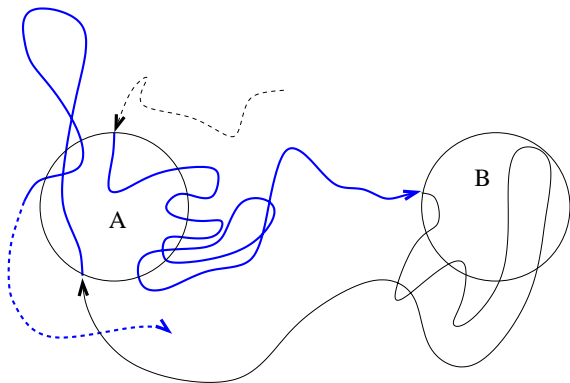
Works in progress:

- Implementations and tests in the NAMD software (collaboration with SANOFI, C. Mayne and I. Teo), and in TRIPOLI (collaboration with CEA)
- Adaptive computation of better and better ξ .
- Analysis of the efficiency as a function of ξ . For optimal choice of ξ , the cost of AMS is (for n large)

$$((\log p)^2 - \log p)$$

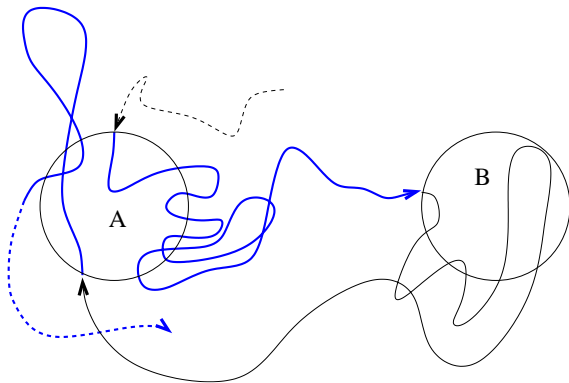
much better than the cost of naive Monte Carlo: $\frac{1-p}{p}$. How does this degrade when ξ departs from the optimal case ?

2- Computing transition times with AMS



Transition time

Let us consider an ergodic stochastic continuous in time process $(X_t)_{t \geq 0}$ in \mathbb{R}^d , and two disjoint subsets $A \subset \mathbb{R}^d$ and $B \subset \mathbb{R}^d$. The objective is to compute the **mean transition time at equilibrium** from A to B , denoted by $\Delta_{A \rightarrow B}$.



Remark: we are also interested in any statistical property of the equilibrium reactive paths from A to B .

Metastability

Examples: Molecular dynamics (A and B are defined in positions space)

- **Langevin** dynamics (M mass matrix, $\gamma > 0$, $\beta = (k_B T)^{-1}$)

$$\begin{cases} dQ_t = M^{-1}P_t dt, \\ dP_t = -\nabla V(Q_t) dt - \gamma M^{-1}P_t dt + \sqrt{2\gamma\beta^{-1}}dW_t, \end{cases}$$

ergodic wrt $\nu(dq) \otimes Z_p^{-1} \exp\left(-\beta \frac{p^t M^{-1} p}{2}\right) dp$ with

$$d\nu = Z^{-1} \exp(-\beta V(q)) dq,$$

where $Z = \int \exp(-\beta V)$.

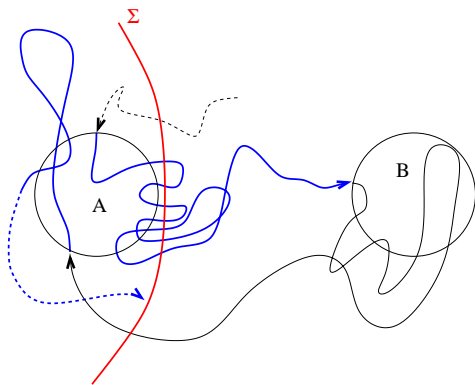
- **over-damped Langevin** dynamics

$$dX_t = -\nabla V(X_t) dt + \sqrt{2\beta^{-1}}dW_t,$$

which is also ergodic wrt ν .

Challenge: A and B are typically **metastable states**, so that observing transitions from A to B is a **rare event**.

From continuous time to discrete time



Let Σ be a co-dimension 1 submanifold in-between A and B . Then, $(Y_n)_{n \geq 0}$ is the sequence of successive intersections of $(X_t)_{t \geq 0}$ with $\mathcal{A} = \partial A$ or $\mathcal{B} = \partial B$, while hitting Σ in-between.

From continuous time to discrete time

More precisely:

$$Y_n = X_{\tau_n}$$

where

$$\tau_n^\Sigma = \inf\{t > \tau_{n-1}, X_t \in \Sigma\}$$

$$\tau_n = \inf\{t > \tau_n^\Sigma, X_t \in \mathcal{A} \cup \mathcal{B}\}.$$

The **Markov chain** $(Y_n)_{n \geq 0}$ is with values in $\mathcal{A} \cup \mathcal{B}$, with kernel:

$\forall x \in \mathcal{A} \cup \mathcal{B}, \forall C \subset \mathcal{A} \cup \mathcal{B},$

$$K(x, C) = \int_{z \in \Sigma} \mathbb{P}^x(X_{\tau_1^\Sigma} \in dz) \mathbb{P}^z(X_{\tau_1} \in C) dz.$$

Reactive entrance distribution

Let us define the successive entrance times in \mathcal{A} and \mathcal{B} [Lu, Nolen, 2013]

[E, Vanden Eijnden, 2006]:

$$T_{k+1}^{\mathcal{A}} = \inf\{n > T_k^{\mathcal{B}}, Y_n \in \mathcal{A}\}$$

$$T_{k+1}^{\mathcal{B}} = \inf\{n > T_{k+1}^{\mathcal{A}}, Y_n \in \mathcal{B}\}.$$

The **reactive entrance distribution in \mathcal{A} at equilibrium** is defined by:

$$\nu_E = \lim_{K \rightarrow \infty} \hat{\nu}_{E,K}$$

where

$$\hat{\nu}_{E,K} = \frac{1}{K} \sum_{k=1}^K \delta_{Y_{T_k^{\mathcal{A}}}}.$$

Remark: ν_E is independant on the choice of Σ and is also the reactive entrance distribution for the original continuous time process.

Back to the mean transition time

The mean transition time at equilibrium is (strong Markov property):

$$\Delta_{A \rightarrow B} = \mathbb{E}^{\nu^E} \left(\sum_{n=0}^{T_B-1} \Delta(Y_n) \right)$$

where

$$T_B = \inf\{n \geq 0, Y_n \in B\}$$

and for all $x \in \mathcal{A}$,

$$\Delta(x) = \mathbb{E}^x(\tau_1).$$

Remark: Notice that

$$\Delta(x) = \mathbb{E}^x(\tau_1 \mathbf{1}_{Y_1 \in \mathcal{A}}) + \mathbb{E}^x(\tau_1 \mathbf{1}_{Y_1 \in B})$$

is the average time of loop from x back to \mathcal{A} when $Y_1 \in \mathcal{A}$ and the average time of a reactive trajectory from x to B when $Y_1 \in B$.

Summary

Objective: Given a discrete-time Markov chain $(Y_n)_{n \geq 0}$ with values in $\mathcal{A} \cup \mathcal{B}$ and a bounded measurable function $f : \mathcal{A} \rightarrow \mathbb{R}$, estimate:

$$\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_{\mathcal{B}}-1} f(Y_n) \right).$$

Two challenges: The sets \mathcal{A} and \mathcal{B} are metastable, so that (i) $T_{\mathcal{B}}$ is very large, and (ii) ν_E is difficult to sample.

Ideas: For (i), use rare event sampling method (forward flux sampling -FFS- or adaptive multilevel splitting -AMS-). For (ii), use the fact that \mathcal{A} is metastable: the process $(Y_n)_{n \geq 0}$ reaches “equilibrium within \mathcal{A} ” (quasi stationary distribution) before transitioning to \mathcal{B} .

Assumptions and notation

Assumptions: In the following, we assume that the Markov chain $(Y_n)_{n \geq 0}$ satisfies the following hypothesis:

[A1] $(Y_n)_{n \geq 0}$ is weak-Feller meaning that $(Kf) \in \mathcal{C}(\mathcal{A} \cup \mathcal{B}, \mathbb{R})$ whenever $f \in \mathcal{C}(\mathcal{A} \cup \mathcal{B}, \mathbb{R})$.

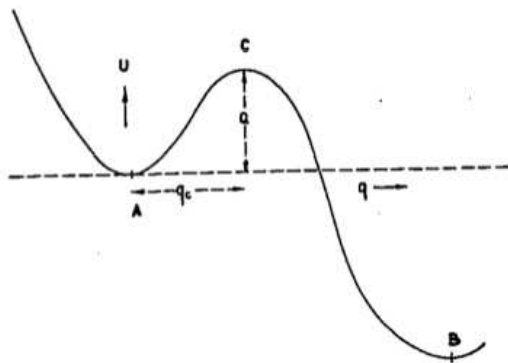
[A2] $(Y_n)_{n \geq 0}$ is positive Harris recurrent, and π_0 denotes its unique stationary probability measure.

[A3] $\pi_0(\mathcal{A}) > 0$ and $\pi_0(\mathcal{B}) > 0$.

All these assumptions are satisfied for the discrete processes built from the Langevin or overdamped Langevin dynamics.

Notation: In the following we use the block-decomposition of the kernel K of the chain $(Y_n)_{n \geq 0}$ over $\mathcal{A} \cup \mathcal{B}$: $K = \begin{bmatrix} K_{\mathcal{A}} & K_{\mathcal{A}\mathcal{B}} \\ K_{\mathcal{B}\mathcal{A}} & K_{\mathcal{B}} \end{bmatrix}$.

The Hill relation



[Kramers, 1940]

The π -return process and the Hill relation

Let π be a probability measure on \mathcal{A} . The π -return process $(Y_n^\pi)_{n \geq 0}$ is the Markov chain with values in \mathcal{A} and transition kernel: $\forall x \in \mathcal{A}, \forall C \subset \mathcal{A}$,

$$K^\pi(x, C) = \mathbb{P}^x(Y_1 \in C, T_B > 1) + \mathbb{P}^x(Y_1 \in B)\pi(C).$$

In words, $(Y_n^\pi)_{n \geq 0}$ is the chain $(Y_n)_{n \geq 0}$ “reset to π ” each time Y_n enters B .

Lemma. $(Y_n^\pi)_{n \geq 0}$ admits a unique stationary distribution, denoted by $R(\pi)$, where

$$R(\pi) = \frac{\pi(\text{Id}_{\mathcal{A}} - K_{\mathcal{A}})^{-1}}{\mathbb{E}^\pi(T_B)}.$$

Remark: Such processes are typically used in MD when people introduce a sink in B and a source in A to create a non-equilibrium flux from A to B [Farkas, 1927] [Kramers, 1940], Weighted Ensemble [Zuckerman, Aristoff], Milestoning [Elber, Vanden Eijnden], TIS [Bolhuis, Van Erp].

The π -return process and the Hill relation

We are now in position to state the **Hill relation** [Hill, 1977] [Aristoff, 2018].

Proposition. For any bounded measurable function $f : \mathcal{A} \rightarrow \mathbb{R}$,

$$\mathbb{E}^\pi \left(\sum_{n=0}^{T_{\mathcal{B}}-1} f(Y_n) \right) = \frac{R(\pi)f}{\mathbb{P}^{R(\pi)}(Y_1 \in \mathcal{B})}.$$

Remark: If $R(\pi)$ is easy to sample, the RHS is typically easier to compute, since it only involves one step of (Y_n) .

Application of the Hill relation to $\pi = \nu_E$

Lemma. The probability measure $R(\nu_E)$ is the stationary distribution π_0 restricted to \mathcal{A} :

$$R(\nu_E) = \frac{\pi_0 \mathbf{1}_{\mathcal{A}}}{\pi_0(\mathcal{A})} =: \pi_{0|\mathcal{A}}.$$

As a consequence,

$$\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_{\mathcal{B}}-1} f(Y_n) \right) = \frac{\pi_{0|\mathcal{A}}(f)}{\mathbb{P}^{\pi_{0|\mathcal{A}}}(Y_1 \in \mathcal{B})}.$$

The Hill relation to compute $\Delta_{A \rightarrow B}$

Back to the mean transition time:

$$\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_B-1} \Delta(Y_n) \right) = \Delta_{Loop}(\pi_{0|\mathcal{A}}) \left(\frac{1}{P_{React}(\pi_{0|\mathcal{A}})} - 1 \right) + \Delta_{React}(\pi_{0|\mathcal{A}})$$

where

- $\Delta_{Loop}(\pi_{0|\mathcal{A}}) = \mathbb{E}^{\pi_{0|\mathcal{A}}}(\tau_1 | Y_1 \in \mathcal{A})$ is the mean time for a loop from $\pi_{0|\mathcal{A}}$ back to \mathcal{A} (computed by brute force Monte Carlo)
- $\Delta_{React}(\pi_{0|\mathcal{A}}) = \mathbb{E}^{\pi_{0|\mathcal{A}}}(\tau_1 | Y_1 \in \mathcal{B})$ is the mean time of a reactive trajectory from $\pi_{0|\mathcal{A}}$ to \mathcal{B} (computed by FFS/AMS)
- $P_{React}(\pi_{0|\mathcal{A}}) = \mathbb{P}^{\pi_{0|\mathcal{A}}}(Y_1 \in \mathcal{B})$ is the probability to get a reactive traj. starting from $\pi_{0|\mathcal{A}}$ (computed by FFS/AMS)

The difficulty is that π_0 and, a fortiori, $\pi_{0|\mathcal{A}}$ are in general unknown and difficult to sample.

Summary

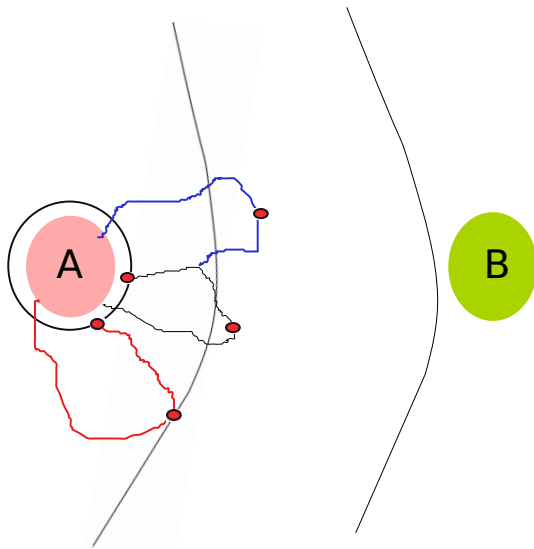
The formula

$$\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_{\mathcal{B}}-1} f(Y_n) \right) = \frac{\pi_{0|\mathcal{A}}(f)}{\mathbb{P}^{\pi_{0|\mathcal{A}}}(Y_1 \in \mathcal{B})}$$

is not practical since $\pi_{0|\mathcal{A}}$ is difficult to sample.

Hope: since \mathcal{A} is metastable, maybe it is not needed to sample ν_E or $\pi_{0|\mathcal{A}}$ since, typically, the process will reach a local equilibrium within \mathcal{A} before going to \mathcal{B} .

A practical algorithm



The quasi-stationary distribution (QSD)

Lemma. Under the assumptions above, the process $(Y_n)_{n \geq 0}$ admits a quasi-stationary distribution (QSD) ν_Q in \mathcal{A} , namely a probability measure ν_Q over \mathcal{A} such that: $\forall C \subset A$,

$$\nu_Q(C) = \mathbb{P}^{\nu_Q}(Y_1 \in C | T_B > 1).$$

In the following, we assume that

[B] $(Y_n)_{n \geq 0}$ admits a **unique** quasi-stationary distribution ν_Q .

Properties of the QSD:

- For any initial condition $x \in \mathcal{A}$, for any $C \subset A$,

$$\lim_{n \rightarrow \infty} \mathbb{P}^x(Y_n \in C | n < T_B) = \nu_Q(C).$$

- The ν_Q -return process admits ν_Q as an invariant distribution:

$$R(\nu_Q) = \nu_Q.$$

The Hill relation applied to $\pi = \nu_Q$

As a consequence

$$\mathbb{E}^{\nu_Q} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right) = \frac{\nu_Q(f)}{\mathbb{P}^{\nu_Q}(Y_1 \in \mathcal{B})}.$$

Remark: Starting from ν_Q , T_B is geometrically distributed, with parameter $\mathbb{P}^{\nu_Q}(Y_1 \in \mathcal{B}) = P_{React}(\nu_Q)$.

Back to the mean transition time [C erou, Guyader, TL, Pommier, 2011]:

$$\mathbb{E}^{\nu_Q} \left(\sum_{n=0}^{T_B-1} \Delta(Y_n) \right) = \Delta_{Loop}(\nu_Q) \left(\frac{1}{P_{React}(\nu_Q)} - 1 \right) + \Delta_{React}(\nu_Q)$$

What did we gain, compared to $\pi = \nu_E$? The probability distribution ν_Q can be sampled by brute force Monte Carlo.

The algorithm to compute $\Delta_{A \rightarrow B}$

In practice:

- Simulate the process $(X_t)_{t \geq 0}$ (or $(Q_t, P_t)_{t \geq 0}$) in a neighborhood of A , registering the successive loops from A to Σ . This gives samples distributed according to ν_Q , and $\Delta_{Loop}(\nu_Q)$.
- Use AMS to simulate reactive trajectories, starting from the QSD ν_Q . This gives an estimate of $P_{React}(\nu_Q)$.

Remark: Typically, one has $P_{React}(\nu_Q) \ll 1$ and $\Delta_{React}(\nu_Q) \ll \frac{\Delta_{Loop}(\nu_Q)}{P_{React}(\nu_Q)}$ so that

$$\mathbb{E}^{\nu_Q} \left(\sum_{n=0}^{T_B-1} \Delta(Y_n) \right) \simeq \frac{\Delta_{Loop}(\nu_Q)}{P_{React}(\nu_Q)}.$$

This is the formula used in FFS to compute transition times [Allen, Valeriani, ten Wolde, 2009].

Error analysis

$$\left| \frac{\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right) - \mathbb{E}^{\nu_Q} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right)}{\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right)} \right| \ll 1?$$

Error analysis

In practice, we thus compute $\mathbb{E}^{\nu_Q} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right)$ instead of the truth $\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right)$.

Objective: Quantify the relative error

$$ERR = \left| \frac{\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right) - \mathbb{E}^{\nu_Q} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right)}{\mathbb{E}^{\nu_E} \left(\sum_{n=0}^{T_B-1} f(Y_n) \right)} \right|.$$

as a function of how large is the transition time wrt the convergence time to the QSD.

Transition time

The time to observe a transition to \mathcal{B} is measured by

$$\frac{1}{\rho^+}$$

where $\rho^+ = \sup_{x \in \mathcal{A}} \mathbb{P}^x(Y_1 \in \mathcal{B})$.

Remark: One obviously has, for any $x \in \mathcal{A}$,

$$\frac{1}{\rho^+} \leq \mathbb{E}^x(T_{\mathcal{B}}).$$

Convergence time to the QSD

The convergence time to the QSD, starting from ν_E is measured by:

$$T_Q^E = \|\nu_E H\|_{TV}$$

where

$$\begin{aligned} Hf(x) &= \mathbb{E}^x \left(\sum_{n=0}^{T_B-1} (f(Y_n) - \nu_Q f) \right) \\ &= \mathbb{E}^x \left(\sum_{n=0}^{\infty} (f(Y_n^{\nu_Q}) - \nu_Q f) \right) \end{aligned}$$

Why can T_Q^E be seen as a convergence time to the QSD?

For any f and for all $n \geq 0$,

$$\mathbb{E}^{\nu_E} (f(Y_n^{\nu_Q}) - \nu_Q f) = \mathbb{E}^{\nu_E} (f(Y_n) - \nu_Q f | T_B > n) \mathbb{P}^{\nu_E}(T_B > n)$$

and thus

$$T_Q^E \leq \sum_{n=0}^{\infty} \|\mathcal{L}^{\nu_E}(Y_n | T_B > n) - \nu_Q\|_{TV}.$$

Error analysis

Proposition. Assume that $\rho^+ T_Q^E < 1$. Then,

$$ERR \leq \frac{\rho^+ T_Q^E}{1 - \rho^+ T_Q^E} \left(1 + \frac{\|f\|_\infty}{|\pi_{0|\mathcal{A}}(f)|} \right).$$

This shows that the error is small if the transition time is large compared to the convergence time to the QSD, i.e.

$$\frac{1}{\rho^+} \gg T_Q^E.$$

Remark: We have checked on examples that the upper bound is sharp in various ways. In particular, one cannot replace ρ^+ by $P_{React}(\nu_Q)$ nor by $P_{React}(\nu_E)$ in the RHS.

Example: the geometrically ergodic case

In the context of the over-damped Langevin dynamics, one can show that: $\exists \alpha > 0, \exists \rho \in (0, 1), \forall x \in \mathcal{A}, \forall n \geq 0,$

$$\|\mathcal{L}^x(Y_n | T_B > n) - \nu_Q\|_{TV} \leq \alpha \rho^n.$$

In this case,

$$T_Q^E \leq \frac{\alpha}{1 - \rho}$$

which goes to zero when $\alpha \rightarrow 0$.

Conclusion (1/2)

We now have a good understanding of the formula which is used by many algorithms (FFS, AMS and the “source and sink methods”: TIS, WE, milestoning) to compute the mean transition time:

- These methods are exact if the process is initialized in the initial state with the correct distribution: the reactive entrance distribution
- The reactive entrance distribution can be replaced by the QSD if A is metastable.

Conclusion (2/2)

Current research directions:

- We analyzed the bias, and not the variance or the efficiency of the whole procedure. This should be possible, at least in simple prototypical cases, and maybe give some hints on good choices for some numerical parameters (position of Σ).
- In practice, it is observed that the initial conditions that indeed yield a transition to B are concentrated on some parts of the boundary ∂A . We are currently working on good sampling methods for these initial conditions.

The problem is typically not to replace ν_E by ν_Q , but to sample ν_Q correctly.

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Simulating dynamics: conclusions (1/2)

There are other mathematical settings to characterize / quantify metastability:

- **Large deviation** techniques [Freidlin, Wentzell, Vanden Eijnden, Weare, Touchette,...] and Onsager-Machlup functionals [Stuart, Pinsky, Theil]
- **Potential theoretic** approaches [Bovier, Schuette, Hartmann,...]
- **Spectral analysis** of the Fokker Planck operator on the whole space and semi-classical analysis [Schuette, Helffer, Nier, Pavliotis]

Simulating dynamics: conclusions (2/2)

There are actually many numerical techniques:

- **Going from state A to state B:**
 - *Local search*: the string method [E, Ren, Vanden-Eijnden], max flux [Skeel], transition path sampling methods [Chandler, Bolhuis, Dellago],
 - *Global search, ensemble of trajectories*: AMS, transition interface sampling [Bolhuis, van Erp], forward flux sampling [Allen, Valeriani, ten Wolde], milestoning techniques [Elber, Schuette, Vanden-Eijnden]
- **Importance sampling approaches on paths**, reweighting [Dupuis, Vanden-Eijnden, Weare, Schuette, Hartmann]
- **Accelerated dynamics techniques** and state to state dynamics [Voter, Perez, Henkelman]
- **Saddle point search techniques** [Mousseau, Henkelman] and **graph exploration**
- **Starting from a long trajectory, extract states**: clustering, Hidden Markov chain [Schuette]

Summary of the algorithms we have discussed

Sampling the canonical distribution:

- Thermodynamic integration: constrained sampling.
- Free energy adaptive biasing methods: importance sampling with an importance function computed on the fly.

Sampling metastable trajectories and rare events:

- Accelerated dynamics algorithms: parallel replica, hyperdynamics, temperature accelerated dynamics.
- Splitting methods: Adaptive Multilevel Splitting.